

Department of Mathematics
Faculty of Mechanical Engineering
Slovak University of Technology in Bratislava

2nd International Conference *APLIMAT 2003*

M | G | ∞ QUEUE SYSTEMS BUSY PERIOD AND LOGISTICS

FERREIRA Manuel Alberto M., (Portugal)

Abstract. In the M $|G| \infty$ queueing systems customers arrive according to a Poisson process at rate λ . Each of them receives immediately after its arrival a service whose length is a positive random variable with distribution function G(.) and mean value α . An important parameter of the system is the traffic intensity $\rho = \lambda \alpha$. The service of a customer is independent of the services of the other customers and of the arrival process. The busy period of a queueing system begins when a customer arrives there, finding it empty, and ends when a customer leaves the system letting it empty. During the busy period there is always at least one customer in the system. Therefore in a queueing system there is a sequence of idle and busy periods.

For these systems with infinite servers the busy period length distribution is difficult to derive, except for a few exceptions. But we present formulas that allow the calculation of some of the busy period length parameters for the $M \mid G \mid \infty$ queueing system.

These results can be applied in logistics. For instance to the failures that occur in the operation of a fleet of aircraft, of shipping or of trucking. The customers are the failures. And its service time is the time that goes from the instant at which they occur till the one at which they are completely repaired. Here a busy period is a period in which there is at least one failure waiting for reparation or being repaired. The formulas referred allow the determination of measures of the system performance.

Key Words: M $\mid G \mid \infty$, Busy Period, Failures

1 Introduction

In the M \mid G \mid ∞ queueing system

- The customers arrive according to a Poisson process at rate λ ,
- Each of them receives a service whose length is a positive random variable with distribution function G(.) and mean value α . So

$$\alpha = \int_{0}^{\infty} [1 - G(t)] dt \quad (1),$$

-There are infinite servers. So when a customer arrives it always finds a server available,

- The service of a customer is independent of the services of the other customers and of the arrival process.

An important parameter is the traffic intensity that we call ρ , being

$$\rho = \lambda \alpha$$
 (2).

It is obvious that in an $M \mid G \mid \infty$ queueing system there are neither losses nor waiting. In fact there is no queueing in the formal sense of the word.

For this systems it is not so important to study the populational process as for other systems with losses or waiting.

Generally we are much more interested in the study of other processes as, for instance, the busy period.

The busy period of a queueing system begins when a customer arrives there, finding it empty, and ends when a customer leaves the system letting it empty. During the busy period there is always at least one customer in the system.

Therefore in a queueing system there is a sequence of idle and busy periods.

We will show in the next section that these concepts can be applied in logistics. Namely to the failures that occur in the operation of a fleet of aircraft, of shipping or of trucking.

The results related to the busy period length of the $M \mid G \mid \infty$ queueing system, that is a random variable, allow the evaluation of performance measures of the fleet. In consequence it is possible to identify ways of improving the performance of the fleet.

We will illustrate the theory with a very simple and short numerical example.

2 Results and Applications

Let us call B the M \mid G \mid ∞ queueing system busy period length.

The mean value of B, whatever is G(.), is given by Takács (1962)

$$E[B] = \frac{e^{\rho} - 1}{\lambda} \qquad (3).$$

But VAR [B], the variance of B, depends largely on the form of B. But Sathe (1985) showed that

$$\lambda^{-2} \max \left[e^{2\rho} + e^{\rho} \rho^2 \gamma_s^2 - 2\rho e^{\rho} - 1; 0 \right] \le VAR[B] \le \lambda^{-2} \left(2e^{\rho} \left(\gamma_s^2 + 1 \right) \left(e^{\rho} - 1 - \rho \right) - \left(e^{\rho} - 1 \right)^2 \right) (4),$$

where γ_s is the variation coefficient of G(.). And, after (4), we can compute easily bounds to SD [B], the standard deviation of B.

Being R(t) the mean number of busy periods that begin in [0,t] (being t=0 the beginning of a busy period) we have, Ferreira (1995),

$$e^{-\rho}(1+\lambda t) \le R(t) \le 1+\lambda t$$
 (5).

Let us call N_B the mean number of the customers served during a busy period in the M $|G| \infty$ queueing systems. We have, Ferreira (2001),

- If G(.) is exponential

$$N_{B}^{M} = e^{\rho} \qquad (6),$$

- For any other distribution function

$$N_{B} \cong \frac{e^{\rho(\gamma_{s}^{2}+1)}(\rho(\gamma_{s}^{2}+1)+1)+\rho(\gamma_{s}^{2}+1)-1}{2\rho(\gamma_{s}^{2}+1)}$$
(7).

These results can be applied in logistics. For instance to the failures that occur in the operation of a fleet of aircraft, of shipping or of trucking. The customers are the failures. And its service time is the time that goes from the instant at which they occur till the one at which they are

completely repaired. For examples of applications of this kind see, for instance, Carrillo (1991). So

- A busy period is a period, in which there is at least one failure waiting for a reparation or being repaired,
- An idle period is a period in which there are no failures present.

Here we gave simple expressions that allow to compute the mean and bounds to the variance of the busy period. And also simple bounds to the mean number of busy periods that begin in a certain length of time. We concluded with expressions to the mean number of failures that occur in a busy period.

These formulas are very simple and of evident application. They only require the knowledge of α , λ , ρ and γ_s that are very easy to compute. The only problem is to test the hypothesis of that the failures occur according to a Poisson process.

Only to conclude note that, calling I (t) the idle period of the $M \mid G \mid \infty$ queueing system distribution function, we have

$$I(t) = 1 - e^{-\lambda t} \tag{8},$$

as it happens with any queue with arrival Poisson process. In this application it gives the probability of that the length of time with no failures is lesser or equal to t.

3 Examples

Suppose a fleet where the failures occur at a rate of 20 per year. So $\lambda = 20/\text{year}$. Suppose too that the mean time to repair a failure is 4 days ($\alpha = 4$ day = (4/365) year). In consequence $\rho \cong 0.22$.

Possibly ρ maybe decreased to 0.11. It is enough to make $\lambda = 10/\text{year}$, for instance buying more vehicles and decreasing, in consequence, the use intensity of each one. Or making $\alpha = 2$ day. For instance increasing the teams that repair the failures.

On other side, if nothing is changed, things can get worse and maybe ρ can increase to 0.44. If we suppose that the repair services times are exponential (a very frequent supposition for this kind of services), $\gamma_s = 1$, and after (3), (4), (5) and (6) we have with t = 1 year (being SD

$$[B] = \sqrt{Var[B]}$$

ρ	E[B]	SD [B] (Lower Bound)	SD [B] (Upper Bound)	R (1) (Lower Bound)	R (1) (Upper Bound)	N_B^M
0.11	2.12 day	2.16 day	2.2 day	18	21	1.12
0.22	4.5 day	4.65 day	4.82 day	16	21	1.25
0.44	10 day	10.72 day	11.46 day	13	21	1.60

And it is possible to conclude that when ρ increases we may have less busy periods in one year, with more failures in each one, of course in mean values.

The mean of the length of the busy period and its dispersion increase with ρ too.

If we suppose now that the repair service times are constant (D = deterministic), $\gamma_s = 0$, and after (3), (4) (in this case the lower bound is equal to the upper bound and so we get the real value of VAR[B]), (5) and (7) we have with t = 1 year

ρ	E[B]	SD [B]	R (1) (Lower Bound)	R (1) (Upper Bound)	N_B^D
0.11	2.12 day	0.41 day	18	21	1.59
0.22	4.5 day	1.22 day	16	21	1.68
0.44	10 day	3.85 day	13	21	1.90

E [B] and the R (1) bounds are the same that in the former case, evidently. The behaviour of the parameters with the increase of ρ is similar to the one of the exponential situation. But now the busy period length dispersion is much lesser and the mean value of failures in each busy period is greater.

4 Conclusion

Of course, in the operation of a fleet, we are interested in big idle periods and in little busy periods. And if these ones occur it is good that they are as rare as possible, with a short number of failures.

Knowing α , λ , ρ and γ_s the manager of the fleet can evaluate the conditions of the operation, namely:

- -The mean length of a period with failures,
- -The length dispersion of a period with failures,
- -The mean number of periods with failures that will occur in a certain length if time,
- -The mean number of failures that occur in a period with failures.

As the expressions depend only on a few parameters and are very simple they show very simple ways to improve the operation, although they may be hard to implement.

References

- 1. Carrillo, M. J. (1991), "Extensions of Palm' s Theorem: A Review", Management Science, Vol. 37, n.º 6, p.739 744.
- 2. **Ferreira**, M. A. M. (1995), "Comportamento transeunte e período de ocupação de sistemas de filas de espera sem espera", PhD Thesis discussed in ISCTE, Supervisor: Prof. Augusto A. Albuquerque.
- 3. **Ferreira**, M. A. M. (2001), "Mean number of the Customers Served During a Busy Period in the M $|G| \infty$ Queueing System", Statistical Review, Vol. 3, INE, 2001.
- 4. **Sathe**, Y. S. (1985), "Improved Bounds for the variance of the busy period of the M | G | ∞ queue", A.A.P., p.913 914.
- 5. **Takács**, L. (1962), "An Introduction to queueing theory", Oxford University Press, New York, 1962.

Current Address

Manuel Alberto Martins Ferreira, Associate Professor, Iscte, Departamento de Métodos Quantitativos, Av. Das Forças Armadas, 1649-026 Lisboa, Telefone: + 351 21 790 30 46, Fax: + 351 21 790 39 64, e-mail: vice.presidente@iscte.pt