

Facing Polyamorous Lives: Translation and Validation of the Attitudes Towards Polyamory Scale in a Portuguese Sample

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Facing Polyamorous Lives: Translation and Validation of the Attitudes Towards Polyamory Scale in a Portuguese Sample

Consensual Non-Monogamies (CNMs) have been receiving the attention of academics, however attitudes towards polyamory are still a new field of research. We aim at studying the reliability and validity of the Attitudes Towards Polyamory (ATP) scale in a Portuguese sample. The present analysis utilized cross-sectional data from 609 volunteers that completed the online survey. After randomly splitting the sample (1:1) for cross-validation purposes, Exploratory (EFA) and Confirmatory factor analysis (CFA) were conducted using Principal Axis Factoring (PAF) and Maximum Likelihood Estimation, respectively. Cronbach's alpha and item-total correlations were used to determine the internal consistency of the scale. The convergent and discriminant validities of the scale were assessed using a correlation matrix. Feasibility and acceptability were examined in terms of missing values, floor and ceiling effects. The ATP proved to be reliable (Cronbach's alpha > .80). In the current study the final structure of the scale, proved through CFA, included 6 items aggregated in a single factor. This final 6-item measure proved to have convergent validity with a measure of modern heterosexism, divergent validity with a measure on self-esteem and concurrent criterion validity with willingness to engage in consensual non-monogamy. Implications for research in polyamory and consensual non-monogamies are discussed.

Keywords: consensual non-monogamies; polyamory; psychometric scale; attitudes; discrimination; Attitudes Towards Polyamory Scale

Introduction

Attitudes tend to aggregate around themselves “beliefs, feelings and behavioral tendencies” (Hogg & Vaughan, 2005, p.150) towards highly salient things in an individual's or society's life. The understanding of attitudes towards a set of identities that are yet targeted for social discrimination might help us understand: 1) how to further the research on attitudes-based discrimination against said groups; 2) how to work with people whose negative attitudes might promote discriminatory behaviors; 3)

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3 how to help those who might feel discriminated to better cope, understand and respond
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5 to those experiences of discrimination. For any of this to be possible, empirical work
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7 demonstrating the (in)existence of positive and negative attitudes is fundamental. To
8
9 achieve this, reliable and valid measures in different languages are necessary. In this
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11 paper, we will look at a measure of attitudes towards polyamory.
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14 15 *The concept*

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17 Polyamory is generally defined as “the assumption that it is possible, valid and
18
19 worthwhile to maintain intimate, sexual, and/or loving relationships with more than one
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21 person” (Haritaworn, Lin, & Klesse, 2006, p.518) at any given time, with the explicit
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23 and informed consent of all those involved, regardless of the existence of sexual
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25 intercourse. There is a definite focus on the emotional and interpersonal aspects of
26
27 relating, with sex being considered as a non-central element, even though sexuality, as a
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29 dimension of human experience and as a practice, keeps being a fundamental marker of
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31 how relationships are made socially intelligible (Mint, 2008) this for the first paragraph
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33 in a section, or to continue after an extract.
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37 Polyamory is considered to be a *relationship orientation*, meaning that many
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39 people feel particularly oriented towards being in polyamorous relationships, just as
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41 others feel oriented towards being in monogamous relationships. Far from being a
42
43 monolithic concept, its vagueness makes it especially permeable to constant social and
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45 subjective reinterpretation, which opens up more potential for discrimination (Barker,
46
47 2005). Overall, polyamory is part of a wider group of sexualized identities and practices
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49 – Consensual Non-Monogamies (CNMs), or “*Open Non-Monogamies*” (Rambukkana,
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51 2015, p. 236) which includes swingers or cuckolders, but not those who are involved in
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53 multiple relationships without the knowledge of all the persons – usually termed
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55 “cheaters” –, for example.
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Discrimination against polyamorous persons

Even though it is still a nascent field within research, there is mounting evidence, internationally, that polyamorous persons are targeted with specific discrimination, in different settings of their private lives, and also at a political level (Cardoso, 2014; Wandrei, 2018). Conley et al. (2012) speak about a “halo effect” surrounding monogamy, since in their studies they found that people would associate more positive characteristics with monogamous persons than with non-monogamous persons, even when said characteristics were unrelated to partnering or sexuality. Other research has demonstrated that clients who are both polyamorous and bisexual require, in therapy settings, specific resources and care from their therapists, but are often advised to ‘return’ to monogamy (Weitzman, 2006). People engaged in CNMs are frequently assumed to have higher sexual infidelity rates by others (Barker, 2005). A recent literature review on several studies around CNMs has shown that several independent researchers demonstrated how “laypeople believe that monogamous relationships are considerably more trusting, committed, passionate, and more sexually satisfying but less likely to involve jealousy than other relational arrangements” (Conley et al., 2012, p. 206). Furthermore, other studies show that there are widely held assumptions about monogamy being a sound strategy to prevent STIs, to improve or maintain relationship satisfaction (Conley et al, 2012). Child-rearing and child welfare is another area where discrimination against polyamory prevails, with monogamous parents being seen as more competent or capable of providing a better child-rearing environment (Conley et al, 2012). This belief is not in line with research that points out that children raised by polyparents have more economic, time and cognitive resources available and that the emotional problems they face come mainly from societal discrimination (Sheff, 2014).

These studies indicate that lay people tend to perceive polyamory as linked to negative outcomes such as emotional instability, negative personal characteristics such

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3 as diminished trustworthiness, and less sexual health, and this may be sustaining
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5 discriminatory attitudes and behaviors towards polyamorous people.
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7 Even though polyamory is a blooming area of research, there is a lack of
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9 measures that take in consideration the experience of polyamory, namely measures that
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11 can assess dimensions linked to knowledge, perception, and potential discrimination
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13 towards polyamory. Understanding attitudes towards minorities and the factors
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15 associated with it is the first step towards a better comprehension of laypeople's
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17 perception of polyamory as well as the understanding of the societal challenges
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19 polyamorous people face.
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22 There is a basis for interrogating the way these phenomena express themselves
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24 in Portugal (Cardoso & Ribeiro, 2016), which justifies our current research. Our review
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26 of the literature demonstrates that there are a few measures at the international level that
27
28 assess attitudes towards relationship diversity, namely polyamory. We found two
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30 measures that assess attitudes towards polyamory, the 12-item *Escala de Atitudes*
31
32 *Frente ao Poliamor* (EAFP) (Freire, 2013) and the 7-item *Attitudes Towards*
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34 *Polyamory* scale (ATP) (Johnson, Giuliano, Herselman, & Hutzler, 2015). Due to its
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36 comparatively more extensive use in international research (with more than a dozen
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38 citations of the original paper in peer-reviewed journals), as well as its briefness, as
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40 compared to Freire's scale, we decided to validate the ATP scale, rather than the EAFP,
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42 with a sample of Portuguese people.
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46 The ATP was developed in the United States in 2015. The development and
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48 validation studies were conducted with 3 distinct samples of men and women: Sample 1
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50 had 100 adult people (38 % women; 62% men; mean age 32.29, SD = 11.18); Sample 2
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52 was developed in two waves with 134 college students (62% women; 37% men, 1%
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54 other; mean age 20.16, SD = 1.77); and Sample 3 had 196 people, (47% women, 52%
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3 Men, 1% other; mean age = 33.28, SD = 12.09). The final version is a unidimensional
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5 measure with 7 items that assesses attitudes towards polyamory on a 7-point Likert scale
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7 from 1 (disagree strongly) to 7 (agree strongly) and has 3 reverse scored items. Higher
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9 total score reflects more positive attitudes towards positive. In these original studies the
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11 measure proved to be valid, explaining 54.8% of variance, with reliability values of
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13 Cronbach's Alpha .86 and temporal stability of $r(128) = .89$. It proved to have
14
15 convergent validity, supported by strong correlations ($r > .50$) with measures of religious
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17 fundamentalism, attitudes towards monogamy, right wing authoritarianism and
18
19 erotophilia and divergent validity with self-esteem and social desirability ($r = .05$ and r
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21 $= .07$).

22 23 24 25 26 ***Aim of the current study***

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28 There are no validated psychometric scales in Portugal that evaluate attitudes towards
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30 polyamory. This study aims to overcome this flaw in the literature and assess the
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32 construct validity and reliability of the ATP scale with a sample of Portuguese people
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34 contributing to expand knowledge on the psychometric behavior of the measure and
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36 allow gathering of further empirical knowledge in the field.

37 38 39 40 41 **Method**

42 43 44 ***Participants***

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46 A total of 609 volunteers (195 male and 414 female), from the general population,
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48 answered the survey. After feasibility and acceptability analysis, cases with missing
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50 values were removed, resulting in a final sample of 519 participants. Of these, 358 were
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52 women (69%) and 161 men (31%) ranging from 18 to 66 years old ($M = 32.80$; $DP =$
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54 10.36), participated in the present study. The sample was predominantly heterosexual
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(69.6%; $n = 346$). Most participants mentioned they were in a monogamous relationship (58.5 %; $n = 299$). The main socio-demographic characteristics of the sample are shown in Table 1. To be included as a participant, respondents needed to be 18 years or older, they had to be living in Portugal, and to have Portuguese as their first language; these conditions were set to try to avoid any problems with question comprehension; respondents also had to identify as men or women (including trans people who identified as such), this condition was set to guarantee that no participants would be excluded due to lack of respondents in other gender categories.

Table 1. Socio-demographic characteristics of the sample (N = 519).

	Men		Women	
	<i>n</i>	%	<i>N</i>	%
<i>Level of Education</i>				
Elementary school	1	0.6	0	0.0
Middle school	4	2.6	6	1.7
High school	41	26.6	87	25
Licenciante	61	39.6	156	44.8
Master	34	22.1	85	24.4
Doctorate	13	8.4	14	4
<i>Sexual orientation</i>				
Heterosexual	109	69	237	69.9
Gay	32	20.3	0	0.0
Lesbian	.0	0.	23	6.8
Bisexual	13	8.2	68	20.1
Queer	4	2.5	7	2.1
Undefined	0	0	4	1.2

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3 The Rosenberg's Self Esteem Scale (RSES) (Rosenberg, 1965) - The RSES
4
5 assess global self-esteem using 10 items that can be answered on a scale from 1 (Strong
6
7 Disagree) to 4 (Strongly Agree). Higher scores indicate higher self-esteem. The measure
8
9 has proven to be valid and reliable in Portuguese samples (Santos & Maia, 2003). In the
10
11 current study it presents a Cronbach's alpha of .92 in the validation sample.
12

13 The Willingness to Engage in Consensual Non-Monogamies Scale (WECNMS)
14
15 (Sizemore & Olmstead, 2017) - It is a 6-item single factor measure. As its name
16
17 indicates the measure assess the willingness to engage in CNM. Respondents answer
18
19 with a scale ranging from 1 (very unwilling) to 7 (very willing) with higher scores
20
21 indicating greater willingness to engage in CNM. The measure has not been validated in
22
23 Portuguese samples. For the current study we followed the procedures for scale
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25 translation and adaptation previously described (page X, line X). In the current study the
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27 measure has proved to be reliable with a Cronbach's alpha of .86. We have, in the
28
29 meantime, become aware that another translation was made by a different Portuguese
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31 team (Lopes & Rodrigues, 2018), and preliminary results, as of yet unpublished, have
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33 demonstrated that the scale is reliable.
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39 ***Procedure***

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41 First, the authors contacted the authors of the original or adapted versions of the
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43 measures in order to ask for authorization to use the measures and pursue the translation
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45 and adaptation when applicable. After the approval by the original author, the research
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47 team developed the translation process following the World Health Organization
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49 protocol of forward-backward translation technique to translating the scale from English
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51 to Portuguese (World Health Organization, 2016). There were two independent
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53 translators of the English version into Portuguese. Some of the authors of the current
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55 study and another person fluent in the two languages evaluated the two versions and
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1
2 merged them based on semantic equivalence. After that, another bilingual person back
3 translated the Portuguese version into English. The author verified the back translation
4 was very similar to the original and therefore a final version was accepted to be
5 validated.
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11 After ethical approval by the XXXXX (Blinded for Review Purposes), the study
12 was set up in a secure server using an online survey platform (LimeSurvey).
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15 Participants were recruited with a snowball-like technique as the URL was disseminated
16 in social and professional networks (e.g. LinkedIn, Facebook) for ten days starting in
17 February 16th, 2017. Potential participants were directed to an informed consent page
18 where information about the authors and aims of study was presented. Information
19 about the voluntary nature of the study as well as confidentiality and non-
20 reimbursement was also provided. The survey included a definition of polyamory
21 before the self-report measures were presented, to minimize the effect of knowledge or
22 disinformation gaps that respondents might have. The definition presented was: “In the
23 context of this study, ‘polyamory’ means the practice, desire or acceptance of being in
24 more than one intimate relationship (sexual and/or amorous, not necessarily romantic)
25 at the same time, with the informed consent of all involved (for instance, someone who
26 has more than one romantic relationship at the same time, and where all people involved
27 know about it and agree to it)”.
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45 ***Procedure***

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47 The psychometric proprieties including feasibility, validity, and reliability of the
48 Portuguese version of the Attitudes Towards Polyamory (ATP) scale were examined.
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50 The feasibility and acceptability were analysed on the entire sample (N = 609) through
51 four indicators: overall response rate, floor effect, ceiling effect (i.e., high endorsement
52 rates at the bottom and top ends of the response scale) and non-response rate. Items
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3 were considered unfeasible when missing values $>.10\%$ and floor and ceiling effects
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5 $<80\%$ (Hilari, Byng, Lamping, & Smith, 2003).
6

7 The validity was assessed using a two-stage process in the exploration and
8
9 validation of the factorial structure of the Portuguese version of the ATP, as
10
11 recommended by Anderson and Gerbing (1988). Cases containing missing values on at
12
13 least one item of the Portuguese version of the ATP were excluded, yielding a final
14
15 sample size of 519 participants (out of 609). The analyses were performed with a split-
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17 sample approach by using the random sample selection procedure (1:1 ratio) in SPSS
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19 version 23.0. This provided a calibration sample ($n = 260$) for identifying and fine-
20
21 tuning the factor structure through exploratory factor analysis (EFA) and a validation
22
23 sample ($n = 259$) for testing the stability of the final model via Confirmatory Factorial
24
25 Analysis (CFA). Following the recommendations of Tabachnick and Fidell (2013), the
26
27 calibration and validation data samples were checked for normality of data distribution
28
29 and outliers, prior to analysis. Both samples were similar in terms of socio-demographic
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31 characteristics (all $ps >.05$). In the first stage of analysis, EFA was conducted using
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33 Principal Axis Factoring (PAF) in order to determine the underlying measurement
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35 model, following the standard factor extraction, rotation, and interpretation phases, as
36
37 recommend by Sakaluk and Short (2017). Prior to EFA, a Parallel Analysis (PA) was
38
39 conducted to determine the number of factors to retain. PA was used in conjunction
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41 with the Guttman-Kaiser criterion (i.e., eigenvalue >1.00). The use of PA relies on the
42
43 fact solid evidence of its accuracy in determining the threshold for significant factors
44
45 and variable loadings when decomposing a correlation matrix (e.g., Franklin, Gibson,
46
47 Robertson, Pohlmann, & Fralish, 1995; R. Ledesma, 2007; Zwick, 2015). PA was
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49 performed with an IBM SPSS MACRO available from O'Connor (2000) and 5000
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51 randomly generated datasets were used. Initially, a factor was retained if the obtained
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3 eigenvalue exceeded the 95th percentile of the random eigenvalue distribution and met
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5 the Guttman-Kaiser criterion. Once the number of factors that met these criteria were
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7 filled, factor solutions were examined. A varimax rotation was applied to transform the
8
9 original principal components produced, to ease interpretation (Hair, Black, Babin, &
10
11 Anderson, 2009). Factor loadings $<.30$ were considered non-substantive, loadings $\geq .30$
12
13 and $<.40$ were considered questionable, and loadings $\geq .40$ were considered substantive
14
15 (Volker et al., 2016). Items with low communalities ($h^2 < 0.3$) were eliminated (Hair et
16
17 al., 2009). Whenever the criteria of factor loading or communality values were not met,
18
19 the item(s) were removed and the EFA was performed again until a final structural
20
21 solution was found. The model derived from the EFA was subsequently evaluated using
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23 CFA at the second stage of validity analysis.
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27 In CFA, the variances of the latent variable were set to unity in order to identify
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29 the structural model and maximum-likelihood estimation procedure was applied.
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31 Multiple criteria were employed to assess the goodness-of-fit of the model (Hair et al.,
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33 2009). The RMSEA was used as the main fit index. RMSEA values are interpreted as
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35 follows: RMSEA smaller than .05 indicates good fit, ranging from .05 to .08 reasonable
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37 fit, .08 to .10 medium fit, and larger than .10 poor fit (Byrne, 2009). Values of CFI,
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39 TLI, and GFI that exceed .9 were interpreted as indicating adequate model fit (Hu &
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41 Bentler, 1998). The modification index (MI) was used for inclusion of additional
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43 parameters. A larger MI (e.g., >50) between two items indicated that those two items
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45 measured the same thing, thus necessitating deletion of one of the items, according to
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47 the parsimony principle (Chang, 2011). The CFA structural model would be modified
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49 until most of the model fit indices meet the criteria. Convergent validity of the items
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51 and factor structure was determined through standardized factor loading (≥ 0.50 were
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53 considered acceptable) and average variance extraction (AVE; ≥ 0.50 was considered
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3 acceptable). Convergent reliability was also assessed through composite reliability
4 (CR), estimated using a covariance structure modeling procedure with nonlinear
5 (CR), estimated using a covariance structure modeling procedure with nonlinear
6 constraints outlined by Raykov (1997) . CR; $\geq .70$ was considered adequate. AVE was
7 manually computed following the guidelines of Hair et al. (2009).
8
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10
11 The reliability of the Portuguese version of the ATP was assessed in both
12 calibration and validation samples. Cronbach's Alpha, inter-item correlation
13 coefficients, corrected item-total correlations and alpha if the item deleted were used to
14 estimate internal consistency reliability. Alpha Cronbach values > 0.7 , item correlation
15 coefficients > 0.20 , and inter-item correlations coefficients < 0.80 and higher than zero
16 were regarded as acceptable. Alpha Cronbach value < 0.5 was regarded as unacceptable
17 (Nunnally, 1978). When the corrected item-total correlation coefficient was < 0.3 or the
18 deletion of which led to an increase of more than 0.1 in Cronbach's coefficient alpha,
19 the item would be removed (Kim & Stoel, 2004; Streiner, 2003).
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31 Finally, convergent, divergent and criterion concurrent validity were assessed
32 using the validation sample. Pearson's correlation coefficient was used to assess the
33 strength of association with related dimensions (Modern Heterosexism); as well as
34 unrelated (Self-Esteem) and a possible outcome (Willingness to Engage in Consensual
35 Non-Monogamy). Pearson r values of 0.10, 0.30 and 0.50 were considered small,
36 medium and large in magnitude, respectively, as recommended by Cohen (1988).
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44 Feasibility, validity (via EFA and Pearson's correlation coefficient) and
45 reliability were tested using IBM SPSS v. 23.0. CFA was performed with IBM SPSS
46 Amos v.20 (SPSS Inc., Chicago, IL, USA). In all statistical procedures, a 5% level of
47 significance was set ($p < .05$).
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Results

Feasibility and Acceptability

Missing data rates for the Portuguese version of the ATP items were low (Table 2), ranging from 2.1 % to 7.4% at an item level. The overall response rate was high (96.5%) with 85.2% participants ($n = 519$) answering all scale items. Both floor and ceiling effects for each item were below 80%, which is acceptable under the recommendations of Hilari et al. (2003). Detailed data related to the feasibility are shown on Table 2.

Table 2. Feasibility and Acceptability results on items of the Portuguese version of the ATP using the entire sample (N = 609).

ATP	<i>N</i>	<i>M</i>	<i>DP</i>	<i>Floor effect</i> (%)	<i>Ceiling effect</i> (%)	<i>Missing data</i> (%)
Item 1	596	5.37	1.86	6.4	43.8	2.1
Item 2	595	5.31	1.77	5.7	37.5	2.3
Item 3	583	5.04	2.00	10.5	37.2	4.3
Item 4	595	5.57	1.67	3.5	43.5	2.3
Item 5	589	5.93	1.55	3.2	54.8	3.3
Item 6	593	5.83	1.60	3.0	54.5	2.6
Item 7	564	3.52	1.87	24.5	9.4	7.4
Overall	587.8	5.22	1.76	8.1	40.1	3.47

A final sample of 519 eligible participants was used for validity and reliability analysis.

Validity

Exploratory Factor Analysis (EFA)

Following the recommendations of Tabachnick and Fidell (2013), calibration sample data ($n = 260$) were checked for normality distribution. Univariate normality was assumed, however, Mahalanobis distance values evidenced 3 multivariate outliers (critical value $\chi^2(7) = 24.32$; $\alpha = .001$), which were removed, resulting on a final calibration sample of 257 participants. PA suggested that the optimal number of factors to retain was only one. This one-factor solution was further examined using PAF. The KMO value of .86 supported the adequacy of the sample. The significance of Bartlett's test of sphericity [$\chi^2(7) = 512.37$; $p < .001$] meant that correlations between items were large enough to conduct an EFA.

After conducting the first EFA, one factor was retained which accounted for nearly 50% of total variance. However, item 7 - "Religious forms of polyamory (such as polygamy) are acceptable" - showed a low communality value (.09) and a factor loading below .4 (.31), which was, therefore, excluded from the further analyses. When this item was removed, the EFA returned again a single factor solution accounting for 53.7% of the explained variance, supporting the unidimensionality of the Portuguese version of the ATP (Table 3).

Table 3. Factor loadings and communalities for each item.

<i>Item</i>		<i>Factor</i>	<i>h²</i>
<i>ITEM 1</i>	Polyamory is harmful to children (R)	.65	.42
<i>ITEM 2</i>	Polyamorous relationships can be successful in the long term	.74	.57

1				
2				
3	<i>ITEM 3</i>	I think that committed relationships with	.74	.54
4				
5		more than two individuals should have		
6				
7		the same legal rights as married couples		
8				
9	<i>ITEM 4</i>	People use polyamorous relationships as	.71	.50
10				
11		a way to cheat on their partners without		
12				
13		consequence (R)		
14				
15	<i>ITEM 5</i>	I would allow my children to spend time	.58	.32
16				
17		with a peer who had polyamorous		
18				
19		parents		
20				
21				
22	<i>ITEM 6</i>	Polyamorous relationships spread STIs	.58	.33
23				
24		(sexually transmitted infections) (R)		
25				
26	<i>Eigenvalue</i>		3.22	
27				
28	<i>Total variance</i>		53.74	
29				
30	<i>explained (%)</i>			

Note: (R) indicates item is reverse coded.

Confirmatory Factor Analysis (CFA)

The validation sample was checked for multivariate outliers via Mahalanobis distance. No multivariate outliers were found (all p_1 and $p_2 > .001$) (Tabachnick & Fidell, 2013). Skewness and kurtosis values were within the recommended range values, indicating a normal distribution of each item (Kline, 2005). The results of the EFA were used as evidence of the underlying one-factor structure (6 items) of the Portuguese version of the ATP and CFA served to confirm this factor structure. The one-factor model showed a good fit TLI = .90, CFI = .94, GFI = .94, however, RMSEA = .122; 90% CI [.09, .16] was unsatisfactory. According to the MI, an improved model fit could be achieved through the addition of a covariance path between e2-e3 (MI = 19.35). Once the error

covariance was added, CFA was reconducted. The new model with correlated errors resulted in better fit measures TLI = .98, CFI = .98, GFI = .98 and RMSEA = .059; 90% CI [.00, .10], indicative of a good fit. All standardized loadings were relatively high, ranging in absolute value from 0.64 to 0.77 and statistically significant, which confirmed the convergent validity of the Portuguese version of the ATP. (Hair et al., 2009). The value of CR was 0.82, indicating good construct reliability. The AVE for the construct was 0.60, suggesting adequate convergence (Hair et al., 2009).

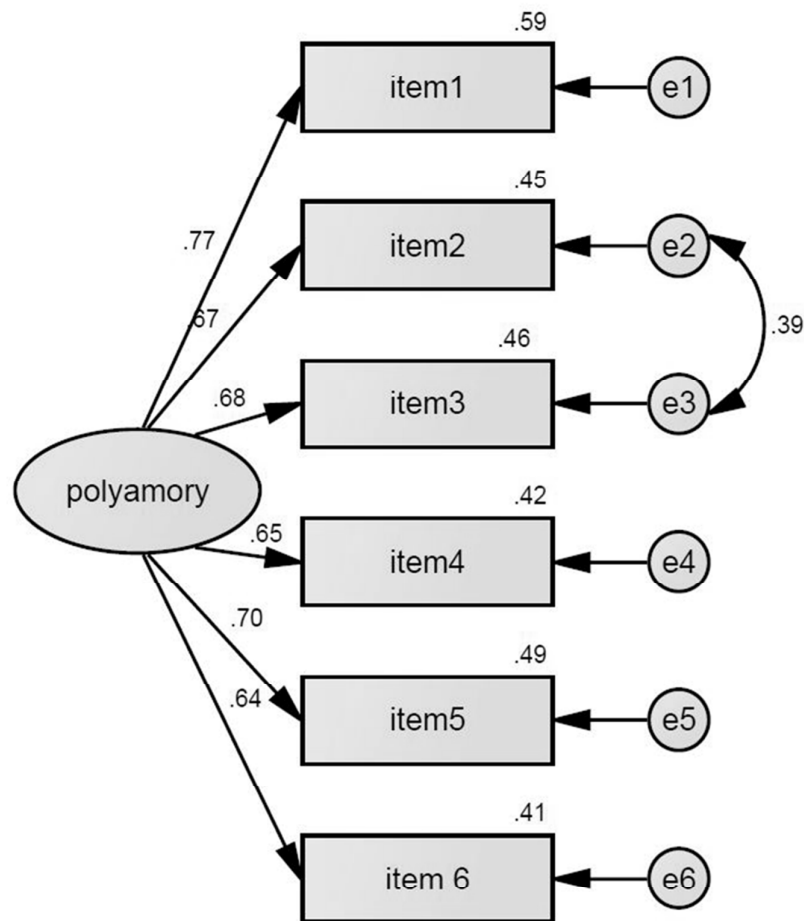


Figure 1. Confirmatory Factor Analysis of the Portuguese version of the Attitudes Toward Polyamory Scale (ATP).

Reliability

The analysis showed a very good internal consistency of the scale in the calibration sample ($\alpha = .83$) and in the validation sample ($\alpha = .85$). Similarly, inter-item correlations and corrected item-total correlations were within the recommended range, indicating homogeneity of the measure (Table 4). Regarding Cronbach's alpha if item deleted, there was no evidence suggesting the removal of any item.

Table 4. Cronbach's alpha coefficients and items correlations for Portuguese version of the Polyamory in the calibration sample and in the validation sample.

	Calibration sample (n=257)		Validation sample (n=259)	
	Corrected item-total correlation	Cronbach's Alpha if Item Deleted	Corrected item-total correlation	Cronbach's Alpha if Item Deleted
<i>ITEM 1</i>	.58	.80	.67	.82
<i>ITEM 2</i>	.66	.78	.67	.82
<i>ITEM 3</i>	.66	.79	.67	.82
<i>ITEM 4</i>	.64	.79	.59	.83
<i>ITEM 5</i>	.52	.81	.63	.83
<i>ITEM 6</i>	.53	.81	.57	.83
<i>Inter-item correlation</i> (<i>M, Range</i>)	44 [.35, .59]		48 [.38, .67]	
<i>Cronbach's Alpha</i>	.84		.85	

Note: M = mean; R = range

Convergent, Divergent and Concurrent Criterion Validity

The associations among variables are presented in Table 5.

Table 5. Pearson's Correlations between attitudes towards Polyamory, self-esteem and willingness to engage in consensual non-monogamies (n=259).

	Modern Heterosexism	Self-Esteem	Willingness to Engage in CNM
Attitudes towards Polyamory	-.61**	.17*	.35**

Note. * $p < .01$; ** $p < .001$

Discussion

The current study aimed to adapt and examine the reliability and validity of a measure of attitudes toward polyamory in a sample of Portuguese adults. We used two subsamples from the same data set and determined that the measure was unidimensional, with a clear structure, and reliable.

The low non-response rate, as well the low floor and ceiling effects observed, suggest all together that Portuguese version of the ATP has good feasibility.

Regarding the results of the EFA, our data is consistent with the original studies of the ATP, with the items loading into a single factor (Johnson et al., 2015). However, in our study the item "Religious forms of polyamory (such as polygamy) are acceptable" was removed as the assumptions for its preservation were not met. There are several aspects that may explain the fact that in our sample this particular item did not prove to load on the single factor. People could have difficulties understanding the concept of polygamy, or could be unfamiliar with the term. The most parsimonious would be that our sample did not recognize the concept of polygamy. Another

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3 explanation can be that due to lack of contact with polygamous religions, participants
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5 could not accurately place their own opinion on the response scale available; according
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7 to data from the Catholic Church, about 89% of Portugal's residents are registered as
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9 Catholic, and other religions that are stereotypically associated with polygamy
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11 constitute only a fraction of the total population. Both explanations may also apply to
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13 the higher non-response rate found in this item. Considering that this item had the lower
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15 mean value (3.52 on a scale range from 1 to 7) it can also be a signal of low tolerance to
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17 religious diversity in our sample, or that respondents chose the central answer in the
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19 scale as a way to signal a 'neutral' response.
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22 The inspection of the mean values of the other items that comprise the measure
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24 reveal that overall participants have a positive attitude towards polyamory. This can
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26 partially be explained by different factors. Firstly, the sample bias. Our sample is young,
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28 highly educated people self-defined with distinct identities, sexual and relationship
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30 orientations. These can be characteristics of people who are more open, aware and
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32 friendly towards sexual diversity, and thus towards non-normative relationship
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34 configurations. Secondly, about 40% of our sample was not in a monogamous situation,
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36 and therefore these results can be indicative of the diversity found in the behaviors of
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38 the respondents themselves. A two-fold final explanation could relate to self-selection
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40 bias: respondents being more willing to complete the survey only when they felt that
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42 their attitudes aligned with the survey's content, and the authors' social and professional
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44 networks might have made the survey more readily available to those with a more
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46 positive outlook on relationship diversity.
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50 The variance explained is within the acceptable range (Streiner, 1994). The CFA
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52 study with the final structured found with EFA showed a good fit. However, in order to
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54 achieve a good fit, a free parameter was needed between the error terms of item 2
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3 (“Polyamorous relationships can be successful in the long term”) and item 3 (“I think
4 that committed relationships with more than two individuals should have the same legal
5 rights as married couples”). In our view this covariance can be theoretically explained
6 by an internalization of a norm that relationships should be long-term, and that long-
7 term relationships are more worthy of institutional protection. As we can see from the
8 phrasing above, the main construct of item 2 is temporality and the main construct of
9 item 3 is legal recognition – but item 3 incorporates the idea of commitment, which is
10 often stereotypically conflated with longer-lasting relationships. Thus, a part of item 3
11 can be understood as overlapping with item 2, and our results bear this out. In addition,
12 the AVE was above 0.50, supporting the validity of the individual indicators in the
13 Portuguese version of the ATP.
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26 Our results based on Cronbach’s Alpha and CR demonstrated that reliability was
27 good. Corrected item-total correlations were similarly high for all items (>.50),
28 indicating homogeneity of the scale. However, the mean inter-item correlation is
29 slightly above the highest recommend value (.40) which indicates that the items are
30 slightly redundant amongst themselves. In fact, there are some authors who claim that
31 values between 0.2 and 0.4 are optimal (Clark & Watson, 1995), while others advocate
32 that a mean inter-item correlation consistently above 0.70, may indicate redundancy
33 (Ponterotto & Ruckdeschel, 2007). Therefore, due the fact that the largest inter-item
34 correlation found in both samples (calibration sample and validation sample) was $r=.67$
35 we considered mean inter item correlations as satisfactory.
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48 The measure is highly related with measure of modern heterosexism, which
49 points towards an integrative approach of understanding discrimination and negative
50 attitudes, linking homophobia to sexism (Pharr, 2002) and, we would argue, also
51 connecting all of this to negative attitudes towards polyamory, since both monogamy
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3 (Pieper & Bauer, 2005) and heterosexuality (Rich, 2007) are part of the normative
4 aspects of sexual and intimate relating in contemporary Western societies (Rubin,
5 2007). Though it is outside the scope of this study, connections between heterosexism
6 and mononormativity should be further investigated.
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11 There is a low association of ATP and self-esteem a result that is different to the
12 one in the original study, where no significant association was found. The magnitude
13 of the association is very weak which confirms that the measure is not relevantly related
14 with constructs that are theoretically unrelated with polyamory.
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20 Additionally, in the current study we explored the possibility that the measure
21 would have concurrent criterion validity with a theoretically relevant future behavior.
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23 Our results support the theoretical hypothesis that more positive attitudes toward
24 polyamory are related to willingness to engage in non-monogamy. This result raises the
25 theoretical possibility that relationship practices can also be shaped by internalized
26 levels of discrimination, and thus gives weight to the importance of combatting
27 discrimination to allow for a more diverse society.
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35 Overall our results support that this version of the ATP is fit to be used in
36 Portuguese. However, this study has limits that cannot be overlooked. Firstly, this is a
37 non-representative sample. Moreover, as stated above, there may be a sampling bias as
38 people more favorable to relationship diversity and with an interest in polyamory may
39 have more will to participate in this study. Furthermore, we did not develop a test-retest
40 analysis that allows to evaluate temporal stability
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48 Nevertheless, this study demonstrates that the Portuguese version of the ATP is
49 a valid, reliable, feasible and well-accepted scale and can be used in future research in
50 polyamory. We have demonstrated that one of the items may be culture sensitive, a
51 result worth exploring in future adaptations of the measure in different cultural contexts.
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3 Furthermore, we have demonstrated that attitudes towards polyamory are strongly
4 associated with attitudes towards other sexual minorities (in the current study, lesbian
5 women and gay men) as well as with the will to engage in consensual non-monogamies.
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7 We have given an additional contribution to the literature in the field by briefly
8 assessing construct validation and reliability of the WECNMS in a Portuguese sample.
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10 Further studies need to be developed in order to better understand the behavior of this
11 measure, especially its temporal evolution, and its connection to other attitudinal
12 constructs and to potentially discriminatory self-reported behaviors.
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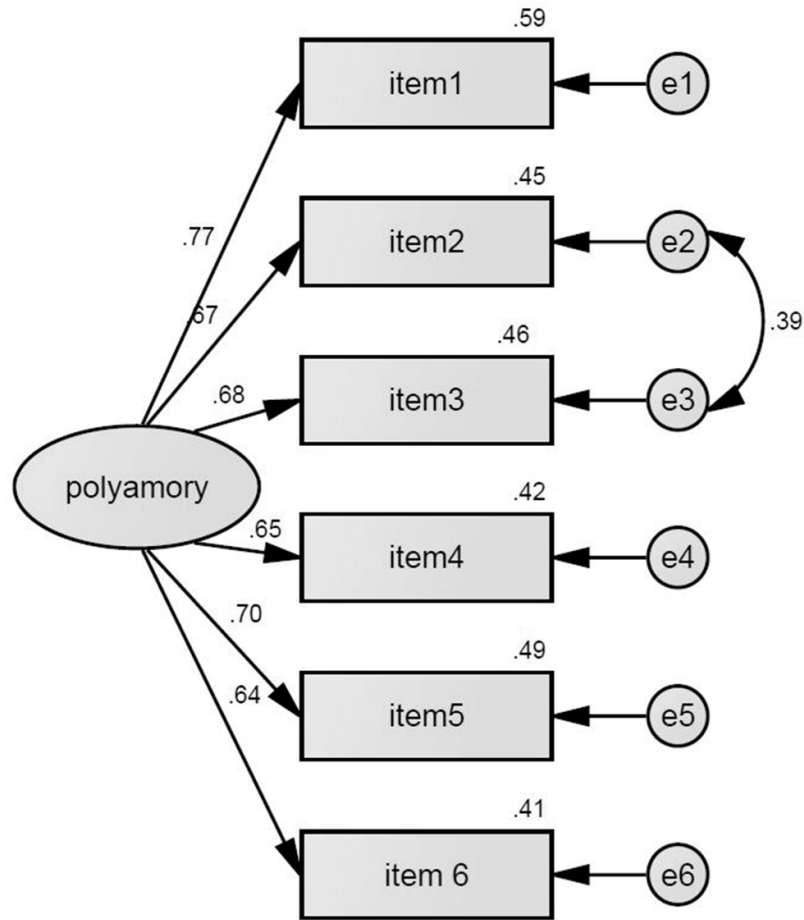


Figure 1. Confirmatory Factor Analysis of the Portuguese version of the Attitudes Toward Polyamory Scale (ATP).