

HOW MUCH DEBT IS TOO MUCH DEBT?
AN EMPIRICAL ANALYSIS ON EMU COUNTRIES

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Dissertation submitted as partial requirement for the conferral of
Master in Economics

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September, 2017

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Abstract

In the aftermath of the global financial crisis, and the subsequent sovereign debt crisis, the relationship between public debt and economic growth was put at the centre of economic policy discussions, specifically regarding the possible existence of a maximum public debt threshold after which economic growth would be severely impaired.

By analysing the short-run relationship between public debt and economic growth for a panel of 10 European Monetary Union (EMU) countries over a period of 22 years, from 1995 to 2016, our results suggest that, by following Reinhart and Rogoff (2010) rationale, the debt-to-GDP ratio threshold would be at 120%. Regarding the direction of causality, our results suggest that public debt causes economic growth in most of the countries analysed. However, by splitting the analysis between the pre-crisis and post-crisis periods, we found evidence that the global financial crisis had an effect in changing the causality in the debt-growth relationship. This analysis also showed that the causal relationship varies across countries. Finally, our threshold analysis provided evidence of a Laffer-curve (inverted U-shape curve) relationship between public debt and economic growth, for which we estimated a maximum public debt threshold around 86% debt-to-GDP ratio.

Considering our results and the current public debt levels, governments should pursue policies that allow the decrease of their level of indebtedness to improve economic performance.

JEL Classification: C33, E62, H63, O47

Keywords: Public debt; economic growth; Granger-causality; public debt threshold effects; panel data analysis; European countries.

Resumo

Na sequência da crise financeira global, e da subsequentemente crise das dívidas soberanas, a relação entre dívida pública e crescimento económico foi colocada no centro da discussão de políticas económicas, nomeadamente quanto à eventual existência de um nível máximo de dívida pública a partir do qual o crescimento económico seria severamente penalizado.

Ao analisar a relação de curto-prazo entre dívida pública e crescimento económico para um painel de 10 países da União Monetária Europeia (UME) durante um período de 22 anos, desde 1995 até 2016, os nossos resultados sugerem que, ao seguir o racional de Reinhart and Rogoff (2010), o rácio dívida-PIB limite seria de 120%. Relativamente à direção de causalidade, os nossos resultados sugerem que dívida pública causa crescimento económico na maioria dos países analisados. No entanto, ao segmentar a análise entre os períodos pré-crise e pós-crise, constatámos que a crise financeira global teve o efeito de alterar a causalidade na relação dívida-crescimento. Esta análise indicou também que a relação de causalidade varia entre países. Por fim, a análise de limites revelou a existência de uma relação curva-de-Laffer (Curva em forma de U invertida) entre dívida pública e crescimento económico, para a qual estimámos um limite de cerca de 86% rácio dívida-PIB.

Atendendo aos nossos resultados e aos níveis atuais de dívida pública, os governos devem conduzir políticas que permitam a redução do seu nível de endividamento de forma a melhorar o desempenho económico.

Classificação JEL: C33, E62, H63, O47

Palavras-Chave: Dívida pública; crescimento económico; causalidade de Granger; efeitos de limites de dívida pública; análise de dados em painel; países Europeus.

Acknowledgements

The love of a family is a life's greatest blessing and I have been blessed by having such a wonderful family. Therefore, I would like to dedicate this master thesis to my mother and father for all the love, support, motivation and advice not only during the development of this research but also throughout my entire life. Thank you for everything, I owe it all to you.

To my supervisor Prof. Doutora Sofia de Sousa Vale for all the support and advice throughout the development of this research. Thank you for all your words of encouragement and guidance throughout this difficult but rewarding journey.

To all my friends that were always supportive and comprehensive of my commitment towards this research. Thank you for all the support, motivation, fun times and for understanding the many moments that I could not be present.

Agradecimentos

O amor de família é a maior bênção da vida e eu fui abençoado ao ter uma família magnífica. Como tal, gostaria de dedicar esta tese de mestrado à minha mãe e ao meu pai por todo o amor, apoio, motivação e conselhos, não só durante o desenvolvimento desta investigação, como também durante toda a minha vida. Obrigado por tudo, eu devo tudo a vocês.

À minha orientadora Prof. Doutora Sofia de Sousa Vale por todo o apoio e conselhos durante o desenvolvimento desta investigação. Obrigado por todas as palavras de alento e orientação durante esta jornada difícil, mas gratificante.

A todos os meus amigos que sempre apoiaram e compreenderam o meu compromisso para com esta investigação. Obrigado por todo o apoio, motivação, tempos divertidos e por compreenderem os vários momentos em que não pude estar presente.

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1 Introduction

Conventional economic theory states that when an economy is depressed governments should intervene by increasing public expenditure to mitigate the negative effects of the downturn phase of the business cycle. Therefore, public debt is seen as a tool for governments to stabilize the economy and smooth business cycles (Puente-Ajovín and Sanso-Navarro, 2015). However, in the aftermath of the global financial crisis that started in 2007/2008, by applying this recipe, with the implementation of stimulus packages and bailing out financial institutions, governments witnessed a sharp increase in their public debt levels (see Figure 1).

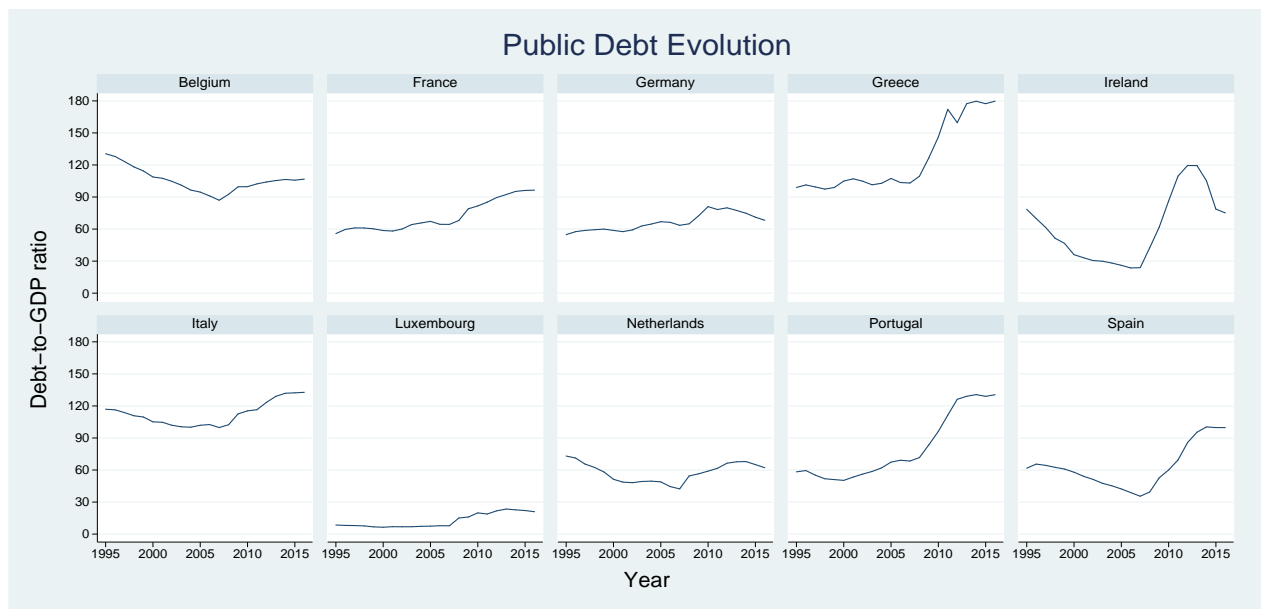


Figure 1: **Public Debt (% of GDP) from 1995 to 2016. Source: AMECO**

With public debt levels presenting an upward trend while the economic context of the time was still unpredictable, the markets and creditors started to show confidence issues about the countries' solvency, *i.e.* the countries' ability to comply with their debt obligations, which was reflected by the higher interest rates demanded by creditors and also through lower credit ratings from the credit rating agencies. Therefore, the risk premium demanded by these economic agents dramatically increased for some European economies, *e.g.* Greece, Ireland, Italy, Portugal and Spain, which witnessed the steepest upward shifts in public debt after 2007 (as seen in Figure 1). This situation made more difficult for these countries to finance their economy with sustainable interest rates, which caused the need for the International Monetary Fund's (IMF) bailouts to enter in the scene, and the global financial crisis evolved into the sovereign debt crisis in Europe. By this time, Reinhart and Rogoff (2010) had

already published their research claiming the existence of a maximum public debt threshold around 90% debt-to-GDP ratio, above which economic growth would be severely impaired, providing empirical support to the conviction that high levels of public debt were damaging to a country's economic performance.

This research intends to shed light on the relationship between public debt and economic growth for a panel of 10 European Monetary Union (EMU) countries from 1995 to 2016. We started by replicating the work of Reinhart and Rogoff (2010) to assess whether the 90% debt-to-GDP ratio threshold holds for our dataset. Then, by following the literature about causality, we analysed the causal relationship between public debt and economic growth for the entire sample, to test whether the assumption of public debt causing economic growth was empirically valid, and also by distinguishing the periods before and after the start of the global financial crisis, to assess whether such event had an effect in changing this specific macroeconomic relationship. Finally, we used methodologies that allowed to endogenously determine public debt thresholds, such as the Hansen (2000) procedure, instead of using *ad-hoc* debt-to-GDP ratio ranges, similarly to Reinhart and Rogoff (2010).

Our replication exercises' results suggest that the 90% debt-to-GDP ratio does not present a turning point for our dataset and, instead, such turning point would be at 120% debt-to-GDP ratio. The causality analysis' results suggest that the causal relationship varies across countries and that the causality direction departed predominantly from public debt to economic growth in most of the EMU countries. Additionally, we found evidence that the global financial crisis had an effect in changing the causality in the debt-growth relationship for some of the countries analysed. Finally, our public debt threshold analysis provided evidence of a Laffer-curve type relationship between public debt and economic growth. This suggests the existence of a maximum public debt threshold, which we estimated to be at 85,67% debt-to-GDP ratio, after which additional public debt increments no longer provide a positive stimulus to economic growth.

This research contributes to the literature by integrating the analysis of three dimensions - Reinhart and Rogoff (2010) replication, causality and threshold effects analyses - of the relationship between public debt and economic growth. Additionally, these analyses were conducted upon a dataset composed by a more homogeneous set of countries within a time frame more closely related with the current economic context.

The remainder of this research is organized as follows. Section 2 provides a literature review of the most relevant and recent investigation concerning the relationship between public debt and economic growth. Section 3 presents the data and methodology used to analyse the different topics in this research. Section 4 provides the results and robustness

checks for each of the topics analysed. Section 5 draws some policy implications and Section 6 concludes this research.

2 Literature Review

The relationship between public debt (or government spending) and economic growth has been a concern for some economists for a long time. The problem of debt was first analysed by using theoretical models. However, treating debt in theoretical models may present some issues because debt can either be trivial or intractable (Cecchetti et al., 2011). In a closed economy model structure, debt would be trivial because the demand for loans would be equal to the supply of loans. Debt would be intractable because, if one considers all the factors that lead an economic agent to choose between debt and equity, the model would end up extremely complex to be treated. As a result, most of the literature follows an empirical approach.

Although the problem of debt appears to be difficult to deal with in theoretical terms, to better understand this issue, one has to be aware of the theoretical views on this matter. As such, Elmendorf and Mankiw (1999) survey, provided insights regarding two different theoretical views on government debt: the "Conventional" view and the "Ricardian Equivalence" view. The "Conventional" view described the ideas held by most economists and politicians while the "Ricardian Equivalence" view described the ideas held by a minority of economists.

According to Elmendorf and Mankiw (1999), the "Conventional" view states that the economy is Keynesian in the short-run. Therefore, if governments create a budget deficit, either by increasing government spending or reducing taxes, the economy would grow due to the positive effect on the aggregate demand. However, the same view states that the economy is "Classical" in the long-run, given that the conditions that make aggregate demand matter in the short-run, specifically sticky prices and wages, are less important in the long-run. By keeping the "Conventional" view idea that the Ricardian Equivalence condition does not hold, Elmendorf and Mankiw (1999) showed that government debt could impair economic growth in the long-run, given that the increase in private savings would not fully compensate the decrease in public savings. This situation would lead to a decrease in the national savings and, consequently, to a decrease in total investment and in the capital stock, therefore leading to a higher marginal product of capital, higher interest rates, lower labour productivity and lower wages. As a result, in the long-run, the increase in government debt leads to a crowding-out effect of private investment which negatively affects output growth.

On the other hand, the "Ricardian Equivalence" view supports the idea that government

spending has no meaningful effects on output growth. According to Elmendorf and Mankiw (1999), this theory relies on the fact that the households are sufficiently forward looking and, consequently, anticipate the future increase in taxes due to the present increase in government spending. As a result, households, even though have available a larger disposable income in the present, would not increase their consumption as a response to a tax cut. Instead, households would save the entire tax cut to fulfil the anticipation of the future tax increase. This means that the decrease in public savings would be totally compensated by the increase in private savings. Therefore, government debt has no effect in changing the economy's path.

From the Elmendorf and Mankiw (1999) survey, it is possible to understand that, theoretically, government debt may have a positive effect in the short-run while in the long-run government debt may have a negative effect on output growth, assuming that the Ricardian Equivalence condition does not hold. Nevertheless, according to Panizza and Presbitero (2013), it is also possible for government debt to present short-run negative effects, if it leads the economy to a situation of uncertainty or negative expectations, either by inflation or financial repression.

These notions point towards the idea that public debt might be a good instrument to motivate growth in the short-run. However, it is an instrument that should be used with caution to not lead the economy to an uncertainty path. Therefore, to avoid such scenario, governments should aim to keep government debt at a sustainable level.

Taking this idea into consideration, the Maastricht Treaty, which established the European Monetary Union, defined the criteria that became known as the "Maastricht Criteria" or the "Convergence Criteria". The criteria of fiscal discipline stated that for a country to be eligible to join the European Monetary Union, it should be capable to maintain its *(i)* deficit below 3% of GDP and *(ii)* the government debt-to-GDP ratio below 60%. If a country failed to comply with these two requirements, it would face an Excessive Deficit Procedure (EDP) to converge to the required levels of the public finances. Exceptionally, countries that surpassed these requirements would not face the EDP, if it was a temporary situation and the ratios were converging to the target at a sufficiently rapid pace. The fiscal discipline criteria of the Maastricht Treaty may lead to the questions of *(i)* how much government debt should be considered as sustainable and *(ii)* if the 60% government debt-to-GDP ratio present a tipping point in the relationship between public debt and economic growth. These questions became more present in the heads of researchers in more recent years, more specifically after the global financial crisis.

As it was already discussed, the "Conventional" view states that the economy is Keynesian in the short-run. Therefore, after the financial crisis, governments increased their spending

to mitigate the negative effects of the crisis. As a consequence of these stimulus packages, some European countries witnessed a sharp increase in their government debt-to-GDP ratios, between 2007 and 2011, which raised concerns around the problem of debt overhang discussed by Krugman (1988), *i.e.* the creditors' belief that a given country would not be able to repay their debt. With such expectation, creditors would demand higher interest rates to provide financing to these economies which would lead their debt level to an unsustainable situation. According to Panizza and Presbitero (2013), the non-linearities and threshold effects in the relationship between public debt and economic growth may arise due to the debt overhang phenomenon.

Departing from this theoretical discussion, Reinhart and Rogoff (2010) conducted an empirical research on the relationship between public debt and economic growth. By analysing the median and average growth rates of 20 advanced economies from 1946 to 2009, the authors found evidence for the existence of a maximum public debt threshold at 90% debt-to-GDP ratio, after which economic growth would be severely impaired.

The conclusions drawn from this research were highly influential among politicians in the aftermath of the financial crisis by providing, according to Herndon et al. (2014), intellectual and empirical support for the application of austerity policies to decrease the high levels of government debt in certain countries. Reinhart and Rogoff (2010) was not only influential among politicians and policy making, but also among many authors by sparking a new branch of research in the economic literature, interested in testing the robustness of the exogenous thresholds defined in Reinhart and Rogoff (2010). That was what Herndon et al. (2014) and Égert (2015a,b) did. Herndon et al. (2014) followed a replication exercise while Égert (2015a,b) not only conducted a replication exercise but also put the dataset used in Reinhart and Rogoff (2010) to formal econometric procedures.

Herndon et al. (2014), by primarily focusing on the Reinhart and Rogoff (2010) sample of 20 advanced economies from 1946 to 2009 and applying the same non-parametric methodology, did not reach the same results as Reinhart and Rogoff (2010). Herndon et al. (2014) presented three reasons for this mismatch of results in the replication exercise, namely the selective exclusion of available data, spreadsheet coding errors and the inappropriate method of weighting statistics in Reinhart and Rogoff (2010) research.

Following a different approach, Égert (2015a,b) applied the Reinhart and Rogoff (2010) dataset not only to a replication exercise, *i.e.* to a non-parametric analysis, but also to formal econometric procedures, to verify the existence of non-linearities and endogenous thresholds in the relationship between public debt and economic growth. Égert (2015a), by applying non-linear threshold models, found evidence for a negative non-linear relationship between

public debt and economic growth and that the negative effects of public debt on growth could emerge at much lower levels than those identified in Reinhart and Rogoff (2010), more specifically between 20% and 60% debt-to-GDP ratios. These results should be considered with caution because, according to the author, they are sensitive to changes in the data. Similarly, Égert (2015b) not only suggested that the non-linear relationship between public debt and economic growth was not very robust, but also that the negative effects of public debt on economic growth could occur at government debt levels as low as 20% of GDP. However, by only taking into consideration the general government debt level¹, the author found evidence for the negative effect to appear at a 50% debt-to-GDP ratio threshold. Additionally, the author conducted a country-specific analysis on the debt-to-GDP ratio thresholds that suggested the existence of a large cross-country heterogeneity.

In summary, the robustness of the results presented in Reinhart and Rogoff (2010) have been challenged by the replication exercises and formal econometric procedures conducted in Herndon et al. (2014) and Égert (2015a,b), which set the 90% debt-to-GDP ratio threshold far from a stylized fact. Nevertheless, the research conducted by Reinhart and Rogoff (2010) has to be acknowledged for motivating the emergence of new branches of research in the economic literature. Besides the branch that was previously discussed, related to the authors' interest in verifying the robustness of Reinhart and Rogoff (2010) results, other researchers focused their interest in analysing the existence of endogenous thresholds and non-linearities, and the direction of causality between public debt and economic growth for a variety of groups of countries and time periods.

2.1 Threshold Effects

In this section, we analyse the literature concerned with testing the existence of endogenous thresholds and non-linearities in the relationship between public debt and economic growth. First, we focus on the authors that analyse these matters by using large panel datasets composed by advanced and emerging economies, such as Woo and Kumar (2015), Afonso and Jalles (2013), Eberhardt and Presbitero (2013), Pescatori et al. (2014) and Eberhardt and Presbitero (2015). Then, we analyse the authors focused on a more restrict set of countries, *e.g.* OECD and Euro Area countries, such as Cecchetti et al. (2011), Checherita-Westphal and Rother (2012), Baum et al. (2013), Afonso and Alves (2014), Bilan and Ihnatov (2015) and Mencinger et al. (2015).

¹Reinhart and Rogoff (2010) used central government debt. However, according to Égert (2015b) it is more relevant to analyse general government debt level (consolidated debt level of the central government, social security and subnational government).

Following the publication of Reinhart and Rogoff (2010), some authors followed the trend in the economic research and started to analyse the relationship between public debt and economic growth to assess whether Reinhart and Rogoff (2010) results were robust to different methodologies, different sets of countries and different time frames of analysis.

Woo and Kumar (2015) analysed a panel of advanced and emerging economies from 1970 to 2007 by using a large variety of estimation methodologies. In their research, Woo and Kumar (2015) found evidence for the existence of a non-linear and negative relationship between public debt and economic growth for debt-to-GDP ratios above 90%. More specifically, that a 10 percentage points (p.p.) increase in the initial debt-to-GDP ratio slowed GDP *per capita* growth by 0,2 p.p. per year. The effect of public debt on economic growth was found to be smaller when only considering advanced economies, where the estimations pointed towards a negative effect of 0,15 p.p. on growth when the initial debt level increased by 10 p.p.. According to the authors, this negative effect reflected the decrease in labour productivity growth due to lower levels of investment² and slower capital stock growth per worker.

Afonso and Jalles (2013), by analysing a panel of 155 countries from 1970 to 2008, also found a negative association between public debt and economic growth, provided that countries with debt-to-GDP ratios above 90% presented smaller growth rates than the countries that were able to keep their debt-to-GDP ratios below 30%. Contrary to Woo and Kumar (2015), this research did not find evidence for non-linearities in the relationship between debt and growth, given that the quadratic debt term was not statistically significant. In terms of public debt thresholds, Afonso and Jalles (2013), by using the Hansen (2000) threshold estimation procedure, derived endogenous thresholds of 58% and 79% for the Euro Area and emerging economies, respectively. The authors also analysed the effects of debt maturity on economic growth. By analysing short-term and long-term debt as a percentage of GDP, it was found evidence for a negative impact on GDP growth of both debt maturities, for the entire sample. However, by analysing short-term and long-term debt as a percentage of total debt, and only taking into account the OECD countries, it was found that longer debt maturities were associated with higher levels of growth. Such evidence suggests that, for these economies, a lower frequency of debt payments would benefit countries with difficulties in accessing the financial markets, thus supporting the current discussion of restructuring public debt in such countries. Additionally, this research supported fiscal consolidation policies given that the inclusion of the budget balance in the estimation process consistently provided positive coefficients, implying that such policies promoted growth in a non-Keynesian way. Regarding the channels of transmission of fiscal consolidation policies, the authors identified a positive effect of such policies on private investment, total factor productivity and capital

²A 10 p.p. increase in the initial debt was associated with a 0,4 p.p. decrease in investment.

stock growth.

Eberhardt and Presbitero (2013) analysed the long-run relationship between public debt and economic growth using a large panel of 105 countries from 1972 to 2009. The authors found evidence for differences in the relationship between public debt and economic growth across countries, but did not find evidence for within-country differences. Also, contrary to much of the literature, this research did not find evidence of non-linearities, given that empirical tests supported a linear specification, nor endogenous thresholds.

Pescatori et al. (2014) applied a novel empirical approach to a large dataset developed by the IMF. Similarly to Eberhardt and Presbitero (2013), the authors did not find evidence for the existence of public debt thresholds. The authors also verified that the negative effects of high levels of debt were only meaningful when only taking the short-run into consideration, given that the association between debt and growth became weaker when analysing for the long-run. Another important contribution in this research was that the relationship between debt and growth was more influenced by the trajectory of the debt-to-GDP ratio than by the debt-to-GDP ratio itself, given the evidence of countries with high but declining debt levels presenting growth paths in line with other countries. Even though the authors did not identify a particular threshold level, they found evidence for the association of large levels of debt with more volatile growth which, according to Pescatori et al. (2014), may be detrimental to a country's economic performance.

Finally, Eberhardt and Presbitero (2015) analysed the long-run relationship between public debt and economic growth for a panel of 118 countries from 1960 to 2012. The authors found evidence, in line with most of the literature, that countries with higher debt-to-GDP ratios tend to present a weaker economic performance. Additionally, the authors support the idea that a common debt threshold for all countries is a fallacy provided that, through their research, it was found evidence for heterogeneity in the relationship between public debt and economic growth across countries.

The results obtained by these authors show a curious pattern. The research published before Herndon et al. (2014) found evidence for a more significant negative effect of public debt on economic growth after the 90% debt-to-GDP ratio. However, the authors that published after the Herndon et al. (2014) critique to Reinhart and Rogoff (2010), either found evidence for debt thresholds far below the 90% debt-to-GDP ratio or did not find evidence for a debt threshold at all. Still, most of the literature based on the analysis of large panel datasets found a common ground by supporting a negative relationship between public debt and economic growth. In terms of the channels through which public debt affects economic growth, there also appears to be a general consensus around private investment,

total factor productivity and capital stock growth.

Now we narrow the analysis of the literature on the authors focused in testing the existence of endogenous thresholds and non-linearities in the relationship between public debt and economic growth for smaller panels of countries.

Cecchetti et al. (2011) by analysing 18 OECD countries from 1980 to 2010 found evidence for a government debt threshold at 85% debt-to-GDP ratio. After this threshold, the authors estimated that a 10 p.p. increase in the debt-to-GDP ratio would reduce trend growth by more than 0,1 p.p.. The authors extended their research by analysing other types of debt, such as corporate and household debt, where their findings suggested thresholds of 90% corporate debt-to-GDP ratio and 85% household debt-to-GDP ratio³.

Checherita-Westphal and Rother (2012) analysed 12 Euro Area countries from 1970 to 2008 and found evidence for a concave relationship (inverted U-shape curve) between public debt and economic growth. This non-linear relationship had a turning point between 90%-100% debt-to-GDP ratios, on average, for the 12 countries in the sample. However, the confidence intervals suggested that debt thresholds could go as low as 70% debt-to-GDP ratio. Similarly to Woo and Kumar (2015), the authors follow different approaches to mitigate endogeneity problems such as the use of 1-year and 5-year forward growth rates, trend GDP growth rates and instrumental variables estimation processes. Regarding the channels through which public debt was found to have a non-linear impact on economic growth, the authors identified private savings and total factor productivity. These channels are in line with those found in Woo and Kumar (2015), given that lower private savings may lead to lower investment levels. With less investment, it would be expected to witness a decrease in the capital stock growth and, consequently, lower productivity levels. As a result, this may end up decreasing the disposable income in the economy and negatively affecting growth.

Baum et al. (2013) analysed 12 Euro Area countries from 1990 to 2010. By using a dynamic threshold panel methodology, the authors were able to analyse the relationship between public debt and economic growth on the short-run, contrary to the previous studies that focus on the long-run relationship. The authors found evidence that the short-run impact of debt on GDP growth was positive, which is in accordance to the economic theory. However, this positive effect tends to zero and loses its statistical significance for debt-to-GDP ratios above 67%. Additionally, for debt-to-GDP ratios above 95%, the authors found evidence that further increases in the debt-to-GDP ratio would lead to a weaker economic performance. The confidence intervals showed that the lower threshold could be as low as 63% debt-to-GDP ratio while the upper threshold could be as high as 100% debt-to-GDP ratio. Baum et al.

³It was acknowledged that the household debt threshold was not estimated with precision.

(2013) also identified pressures in the long-term interest rate when public debt-to-GDP ratio surpassed 70%, hinting for a possible crowding-out effect of private investment at this level. Similarly to Afonso and Jalles (2013), this research supports fiscal consolidation policies for highly indebted economies because, according to Baum et al. (2013), the positive short-run effect of additional debt on growth may decrease drastically and even become negative for high initial debt levels.

Afonso and Alves (2014) analysed the effect of public debt on economic growth for 14 European countries from 1970 to 2012. In line with the previous works, Afonso and Alves (2014) provided evidence for a negative relationship between public debt and economic growth, both in the short-run and long-run. Additionally, when the effects of debt-to-GDP ratio and debt service were compared, the authors found evidence that the debt service variables had a more detrimental effect on economic growth. Regarding the possible existence of debt thresholds and non-linearities, the authors provided evidence for a Laffer-curve relationship between debt and growth with a threshold of 75% and 74% for annual and 5-year average growth rates, respectively.

Bilan and Ihnatov (2015) analysed 33 European countries from 1990 to 2011. The results provided evidence for the existence of a non-linear relationship between public debt and economic growth, *i.e.* a Laffer-curve relationship, with a maximum debt threshold at 94% debt-to-GDP ratio, for the whole sample. However, when the sample was split between developing and developed European countries, the authors found evidence that a common debt threshold for all countries was a fallacy. In this regard, the authors identified two different ranges of thresholds, more specifically a 40%-50% debt-to-GDP ratio for the set of developing European countries and a 70%-80% debt-to-GDP ratio for the set of developed European countries. According to Bilan and Ihnatov (2015), debt effects, after this threshold, were expected to be negative due to higher interest rates, fear of debt unsustainability and of fiscal consolidation measures.

Finally, Mencinger et al. (2015) analysed the short-run relationship between public debt and economic growth in a panel of 36 countries⁴. The time period of analysis was not the same for all countries in the sample. More specifically, the analysis on advanced economies covered the period from 1980 to 2010 while for emerging economies it was only covered the period from 1995 to 2010. In order to study the debt-growth relationship, the authors used a generalized economic growth model augmented with a debt variable and, through their analysis, they found evidence for the existence of a non-linear relationship, *i.e.* a Laffer-curve relationship. The authors found evidence that advanced economies tend to have less

⁴31 OECD countries and the remainder 5 were European (non-OECD) countries.

restrictive thresholds given that, for this panel, the threshold for advanced economies ranged between 90% and 94% debt-to-GDP ratios while for emerging economies the threshold ranged between 44% and 45% debt-to-GDP ratios.

This set of literature is also in line with most of the existent literature, by supporting a negative relationship between public debt and economic growth. In terms of debt thresholds, the evidence point towards a debt threshold for the whole sample of countries that, on average, ranges between 63%-100% debt-to-GDP ratios. However, when distinguishing between advanced and emerging economies in these sets of countries, a common evidence found in the literature was that emerging economies faced more restrict debt thresholds than advanced economies. More specifically, emerging economies present a tipping point in the debt-growth relationship between 40%-50% debt-to-GDP ratios while the one for advanced economies ranges between 70%-94% debt-to-GDP ratios. Regarding the channels through which public debt affects economic growth, this set of literature point towards private savings and total factor productivity, in line with those identified in the first part. Additionally, Bilan and Ihnatov (2015) identified other channels such as higher interest rates and the fear of debt unsustainability which would lead to the application of fiscal consolidation policies.

2.2 Causality

Another important branch in the literature is related with causality. All the previous empirical works analysed departed from the assumption that the causality direction went from public debt to economic growth. Most of the literature support this assumption by referring to the negative correlation between public debt and economic growth. However, correlation does not necessarily imply a causal relationship between two variables. Therefore, discussion on this assumption arose given that it is also reasonable to assume that the causality direction might also go from economic growth to public debt. A clear example of such idea is related to the automatic stabilizers (Panizza and Presbitero, 2014; Gómez-Puig and Sosvilla-Rivero, 2015). When the economy is depressed, the unemployment level tends to increase, the level of consumption is lower and, consequently, governments witness lower levels of revenue. In such scenario and in countries with large welfare systems, the automatic stabilizers' effects kick in, *i.e.* unemployment and other social security benefits are needed by more people, and governments need to accommodate such needs. Therefore, if governments face lower levels of revenue when the economy is depressed and the financing needs increase due to the automatic stabilizers, the only path available for governments to accommodate such expenses might be through public debt.

This branch of research is more recent and the literature on this matter is not as large as

the literature focused on public debt thresholds. Nevertheless, some authors already turned their attention on this particular topic by employing formal econometric procedures to assess the causality direction on a variety of groups of countries and time periods.

Panizza and Presbitero (2014) analysed a sample of OECD countries⁵ by using an instrumental variable approach, capturing the valuation effects caused by the interaction between foreign currency debt and exchange rate volatility. The authors did not find evidence supporting the causality direction from public debt to economic growth. Also, the authors did not find evidence that high public debt levels negatively affect future growth in advanced economies. Nonetheless, such results may be, according to Panizza and Presbitero (2014), due to the fact the countries' public debt levels were still below the country-specific debt thresholds and also due to the fact that, in the time period of analysis, the countries in the sample could use their own central banks, *e.g.* through debt monetization or competitive devaluation policies, to solve their debt problems. Given their results, they argued that the negative relationship between public debt and economic growth, found in most of the literature, should not be used as a justification for fiscal consolidation policies, as there was not found evidence for the causal effect to depart from public debt to economic growth.

Donayre and Taivan (2017) analysed the country-specific causal relationship between public debt and economic growth in a sample of 20 OECD countries from 1970 to 2009. The authors' research was based on canonical cointegrating regressions, which allowed for the possibility of stochastic cointegrating vectors, and then used Granger-causality and VAR econometric tests to make inference about the direction of causality. According to the authors, this methodology allowed them to address the issue of the possible existence of a dynamic relationship between public debt and economic growth that was not addressed by most of the empirical research on the debt-growth nexus. Through this research, Donayre and Taivan (2017) found evidence that modern welfare states face periods of low growth following increases in the debt-to-GDP ratio. However, when analysing more traditional welfare states and welfare states with larger governments, *i.e.* with larger levels of public expenditure, the authors found evidence supporting a bi-directional causality. Such evidence point towards the importance of a country-specific analysis on the debt-growth nexus and that homogeneous fiscal consolidation policies throughout the Euro Area may not be the appropriate route to pursue to achieve more growth. Therefore, according to Donayre and Taivan (2017), it cannot be inferred that larger levels of public debt will severely impair growth in all countries.

Puente-Ajovín and Sanso-Navarro (2015), following a Granger-causality approach, analysed a sample of 16 OECD countries from 1980 to 2009. The authors not only consid-

⁵Provided in Cecchetti et al. (2011).

ered government debt but also household and non-financial corporate debt. To control for the cross-country heterogeneity and cross-sectional dependence, a panel bootstrap Granger-causality test was applied. From this analysis, the authors found evidence that public debt did not Granger-cause economic growth and also that non-financial private debt, *e.g.* household debt, did not Granger-cause economic growth. This research's results pointed towards a causality direction from economic growth to public debt. Therefore, such results did not provide intellectual support for the application of fiscal consolidation policies which find justification on the assumption that public debt causes economic growth.

Gómez-Puig and Sosvilla-Rivero (2015) analysed the bi-directional causal relationship between public debt and economic growth in a sample of 11 European Monetary Union countries from 1980 to 2013. The authors followed a Granger-causality approach and endogenous breakpoint tests to address the possible heterogeneity in the bi-directional causality. According to Gómez-Puig and Sosvilla-Rivero (2015), such methodology not only considers the cross-country differences but also the differences over time, addressing the dynamic nature of the relationship between public debt and economic growth, in line with Donayre and Taivan (2017). Similarly to Panizza and Presbitero (2014) and Puente-Ajovín and Sanso-Navarro (2015), this research did not provide evidence for a negative causation between public debt and economic growth up to 2009. After detecting an endogenous breakpoint between 2007 and 2009, the authors found evidence for a negative Granger-causality effect between public debt and economic growth for some of the countries analysed.

Finally, Ferreira (2016) analysed a sample of 28 European Union countries from 2001 to 2012 and, to study the period after the financial crisis, from 2007 to 2012. Following a Granger-causality approach, the author analysed the causal relationship between three different debt categories - public debt, foreign debt and private debt - and economic growth. Through this approach, Ferreira (2016) found evidence for a statistically significant bi-directional Granger-causality relationship between public debt and economic growth, at least in the short-run. The results also provided evidence for a negative and statistically stronger causal relationship from economic growth to public debt in both panels, *i.e.* for the 2001-2012 panel and for the post-crisis panel (2007-2012). Regarding the other types of debt, the results obtained were not statistically strong. Nevertheless, the estimation results point towards a positive bi-directional causal relationship between foreign debt and economic growth. Regarding private debt it was found evidence for a bi-directional causal relationship, more specifically, that private debt has a negative impact on economic growth and economic growth positively affects private debt.

The conclusions drawn by this set of empirical research do not provide consensual results. Some authors either support both directions of causality or no causal relationship at all

between public debt and economic growth.

In summary, even though the literature found a common ground in supporting a negative relationship between public debt and economic growth, it is possible to conclude that there is no common agreement regarding the other main topics of discussion related with the relationship between public debt and economic growth, such as public debt thresholds and causality direction.

3 Data and Methodology

3.1 Data

We used annual data to analyse the relationship between public debt and economic growth for a panel of 10 European Monetary Union (EMU) countries that signed the Maastricht treaty in 1992 and have the Euro as their official currency. Within this set of countries we have Belgium, France, Germany, Greece, Ireland, Italy, Luxembourg, Netherlands, Portugal and Spain. All the variables used in this research were collected from the European Commission's Annual Macro-Economic database (AMECO) from 1995 to 2016, providing 22 observations per country and giving a total of 220 observations for our dataset. The dependent variable is the real Gross Domestic Product (GDP) annual growth rate and the independent variable of interest is public debt (as % of GDP), *i.e.* the debt-to-GDP ratio. In terms of control variables, we used those that have been consistently used throughout the debt-growth literature, such as the logarithm of real GDP *per capita*, trade openness (as % of GDP), population growth, capital stock (as % of GDP), budget balance (as % of GDP) and debt service (as % of GDP).

3.1.1 Structural Considerations

We used several criteria to select the countries and time-frame under analysis. First, to avoid mismeasurement issues and the usage of different methodologies to develop the statistics, we collected our data from a single database (AMECO). Second, the 10 EMU countries that signed the Maastricht treaty were chosen to mitigate country heterogeneity in terms of policy and economic tools available, *i.e.* after signing this treaty these countries could no longer take advantage of isolated monetary policies, such as debt monetization or competitive devaluations to boost growth. Finally, the time-frame chosen to develop our research, from 1995 to 2016, was based on four factors, more specifically the *i)* ESA 2010 statistical framework for which the data available on the AMECO database only goes back to 1995, *ii)* the data

is more closely related with the current economic environment, *iii*) this shorter time-frame of analysis allows to avoid possible sources of bias that large historical datasets cannot, such as, for example, war periods, different political regimes and country-specific monetary policy and *iv*) this period covers the process of the EMU preparation and implementation.

3.2 Methodology

In this research, we analysed three different topics regarding the relationship between public debt and economic growth. We started by replicating the work of Reinhart and Rogoff (2010) to assess whether the 90% debt-to-GDP ratio, presented as a tipping point in their research, holds with our data. Then, we analysed the causal relationship between public debt and economic growth, given that most of the literature departs from the assumption of public debt causing economic growth without clearly testing such hypothesis. Finally, we employed different methodologies that allowed to endogenously estimate a maximum public debt threshold in a panel framework.

3.2.1 Reinhart and Rogoff (2010) Replication

As it was previously referred, by analysing 20 advanced countries from 1946 to 2009, Reinhart and Rogoff (2010) concluded with their research that there was a negative correlation between public debt and economic growth with a maximum public debt threshold around 90% debt-to-GDP ratio.

To verify if this threshold holds in our sample⁶, we conducted a replication exercise⁷, while avoiding the methodological mistakes pointed out by Herndon et al. (2014), that consisted in analysing the average and median growth rates of the countries in our dataset within the debt-to-GDP ratio ranges defined in Reinhart and Rogoff (2010). Then, by pursuing the same methodology, we extended the analysis by including two additional debt-to-GDP ratio ranges, the 90%-120% and +120%.

3.2.2 Causality

Most of the literature supports their research on the assumption that public debt causes economic growth, by referring to the negative correlation that exists between both variables. However, given that correlation does not necessarily imply a causal relationship, our definition

⁶Belgium, France, Germany, Greece, Ireland, Italy, Luxembourg, Netherlands, Portugal and Spain from 1995 to 2016.

⁷Using the Microsoft Office Excel programme.

of causality follows the one of Granger (1969), *i.e.* that a variable X is said to cause Y if by including the past information of the variable X we can better predict the current value of Y (Ferreira, 2016).

Therefore, to analyse the causal relationship between public debt and economic growth, we followed Ferreira (2016) and Gómez-Puig and Sosvilla-Rivero (2015) in terms of the model structure, and used the Dumitrescu and Hurlin (2012)⁸ Granger non-causality test to draw conclusions regarding the causal relationship between these two variables in a panel framework. Furthermore, we not only analysed the full sample causality (1995-2016) but also extended the analysis by studying the causal relationship between public debt and economic growth for the pre-crisis (1995-2006) and post-crisis (2007-2016) periods.

The Dumitrescu and Hurlin (2012) Granger non-causality test allows to assess whether variable X does not Granger-cause variable Y or if variable X Granger-cause variable Y for at least one country, the null and alternative hypotheses, respectively.

$$\begin{cases} H0 : X \text{ does not Granger-cause } Y \\ H1 : X \text{ Granger-cause } Y \text{ for at least one country} \end{cases}$$

As one can clearly verify, if we do not reject the null hypothesis the conclusion is straightforward. However, if we reject the null, we conclude that variable X Granger-cause variable Y for at least one country. This is a drawback of the test because it does not allow to know for which country or countries it is possible to verify the causal relationship under analysis. Therefore, it was needed to conduct country-specific regressions to identify the country or countries for which the causal relationship under analysis was verified⁹.

By using STATA, we conducted the referred test on the following equations:

$$Y_{i,t} = \sum_{k=1}^K \beta_Y Y_{i,t-k} + \sum_{k=1}^K \beta_X X_{i,t-k} + \mu_{i,t} \quad (1)$$

$$X_{i,t} = \sum_{k=1}^K \beta_Y Y_{i,t-k} + \sum_{k=1}^K \beta_X X_{i,t-k} + \mu_{i,t} \quad (2)$$

where Y = real GDP annual growth rate; X = debt-to-GDP ratio annual growth rate; $i = 1, \dots, N$ countries; $t = 1, \dots, T$ years; $k = 1, \dots, K$ lags and μ is the error term.

Additionally, we conducted several robustness checks, based on Ferreira (2016) method-

⁸This test's routine is available on STATA with the name "**xtgcause**".

⁹The routine "**xtgcause**" available for STATA, also conducts country-specific regressions.

ology, by regressing equations (1) and (2) with other estimators, such as the robust Ordinary Least Squares (OLS) estimator, Fixed Effects (FE) estimator to account for country heterogeneity and the Generalized Method of Moments (GMM) estimator to account for potential endogeneity. These regressions allow to test if the current value of variable Y significantly depends on the past information of variable X and if the current value of variable X significantly depends on the past information of variable Y. Then, by following the causality definition of Granger (1969), these regressions allow to infer about the direction of causality between variables X and Y. Specifically, if lagged X is statistically significant in explaining Y, X is said to Granger-cause Y, and if lagged Y is statistically significant in explaining X, Y is said to Granger-cause X.

3.2.3 Public Debt Thresholds

As our last topic of interest, we analysed the short-run relationship between public debt and economic growth, where we also included a squared debt-to-GDP ratio term to test for a Laffer-curve type relationship between these two variables. Our model specification was the following:

$$Y_{i,t} = \beta_0 + \beta_1 X_{i,t} + \beta_2 X_{i,t}^2 + \sum \beta_j Z_{i,t} + \mu_{i,t} \quad (3)$$

where Y = real GDP annual growth rate; X = debt-to-GDP ratio; Z = set of control variables consistently used throughout the literature; $i = 1, \dots, N$ countries; $t = 1, \dots, T$ years and μ is the error term.

To estimate the maximum public debt threshold of equation (3), we used the Hansen (2000)¹⁰ endogenous threshold estimation procedure for which our threshold variable was the debt-to-GDP ratio. Additionally, we complemented this analysis with several robustness checks, by regressing equation (3) with other estimators, such as the Ordinary Least Squares (OLS) estimator, Least Squares Dummy Variables (LSDV) and Fixed Effects (FE) estimators to account for country heterogeneity, and Generalized Method of Moments (GMM) estimator to account for potential endogeneity. After estimating a statistically significant model for each of the referred estimators, we estimated the maximum public debt threshold following a partial derivative approach, similarly to Bilan and Ihnatov (2015) and Afonso and Alves (2014):

¹⁰This test's routine for STATA is available on http://www.ssc.wisc.edu/~bhansen/progs/ecnmt_00.html.

$$\frac{\partial Y_{i,t}}{\partial X_{i,t}} = \beta_1 + 2\beta_2 X_{i,t} = 0 \quad (4)$$

$$\Leftrightarrow X_{i,t} = \frac{-\beta_1}{2\beta_2} \quad (5)$$

4 Results

In this section we provide the results for each of the topics under analysis regarding the relationship between public debt and economic growth. First, we analyse our replication exercise of Reinhart and Rogoff (2010), then the causal relationship between public debt and economic growth and finally the public debt thresholds. To conclude this section we provide a summary of the main findings from this research.

4.1 Reinhart and Rogoff (2010) Replication

We conducted this replication exercise to assess whether the results and conclusions drawn in Reinhart and Rogoff (2010) research hold with our dataset. An important aspect of this replication exercise was that we addressed the methodological mistakes that were exposed by Herndon et al. (2014).

Even though our dataset differs significantly from the one used in Herndon et al. (2014) and Égert (2015a,b), both in terms of countries and time-frame covered, our main conclusions from the replication exercise are quite similar. First, we found evidence supporting the existence of a negative correlation between public debt and economic growth, as supported by most of the literature, within our dataset. Second, we did not find evidence of average negative economic growth for debt-to-GDP ratios above 90% (see Figure 2).

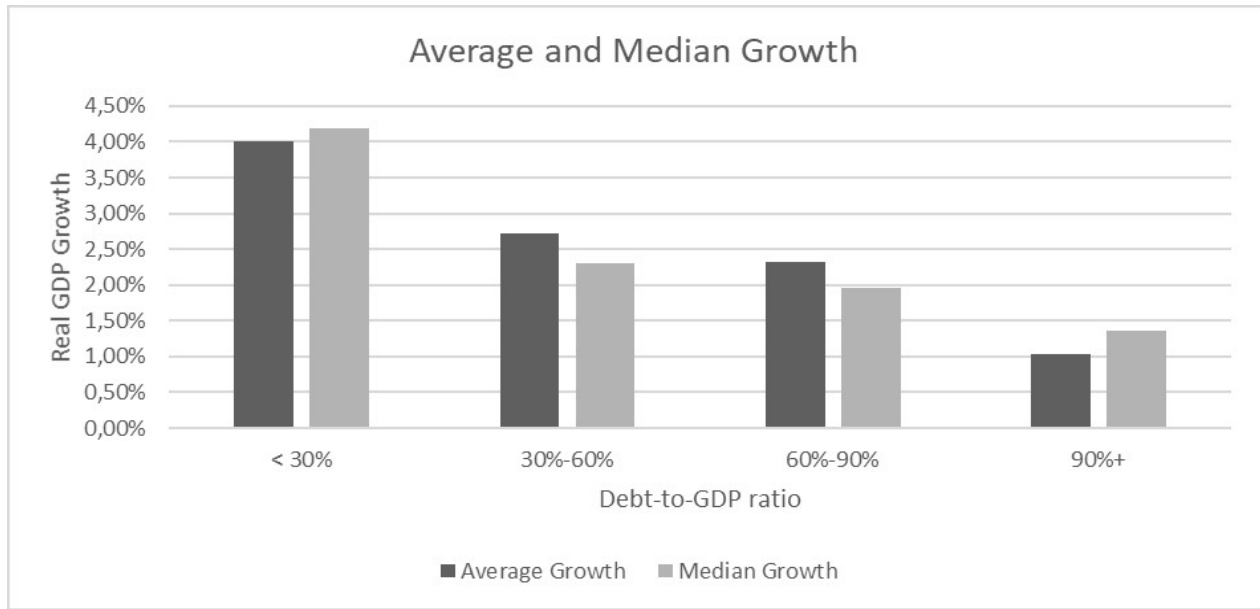


Figure 2: Reinhart and Rogoff (2010) replication results. Source: Authors' calculations

Given that our previous results provided evidence of a negative correlation between public debt and economic growth, we enhanced the replication exercise, by adding two additional debt-to-GDP ratio ranges, to assess whether there was in our data a larger debt-to-GDP ratio range for which we could find evidence of average negative growth.

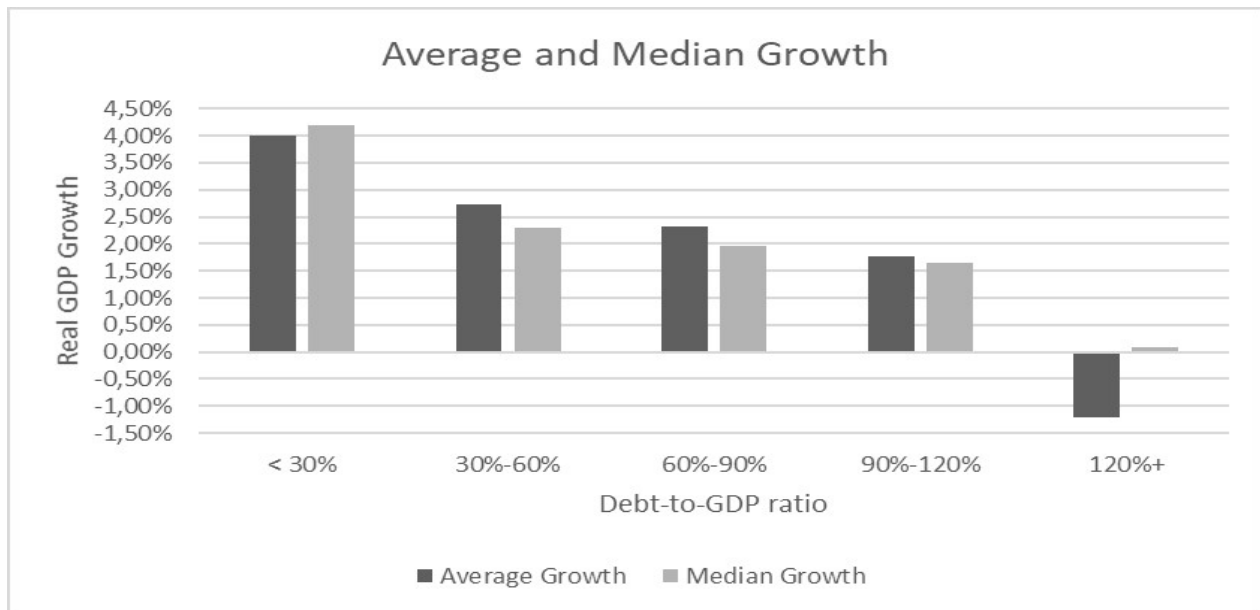


Figure 3: Reinhart and Rogoff (2010) enhanced replication results. Source: Authors' calculations

By including these two additional debt-to-GDP ratio ranges, the 90%-120% and +120%, we still verified the negative correlation between the two variables and we also found evidence of average negative economic growth for debt-to-GDP ratios above 120% (see Figure 3). Therefore, by following Reinhart and Rogoff (2010) rationale, we can claim with our research that the 120% debt-to-GDP ratio presents a turning point in the relationship between public debt and economic growth for the countries analysed in our sample. This result was largely influenced by Greece's poor economic performance during the years within the +120% debt-to-GDP ratio range, specifically from 2009 to 2016.

Nevertheless, we were not tempted to make such claim based on a methodological approach with such rigid assumptions. More specifically, Reinhart and Rogoff (2010) methodology departs from the assumption that the causality direction in the debt-growth relationship goes from public debt to economic growth without testing such hypothesis. Additionally, the thresholds presented by their research appear in the form of *ad-hoc* debt-to-GDP ratio ranges. Therefore, to provide more meaningful results, we conducted further analyses to assess the causality direction between public debt and economic growth and also applied different methodologies to endogenously determine public debt thresholds.

4.2 Causality

Most of the literature on the debt-growth nexus departs from the assumption of public debt causing economic growth, by referring to the negative correlation between both variables, without clearly testing such hypothesis. In this section, we not only analysed the causal relationship for the whole sample period but also for the pre-crisis and post-crisis periods, to assess whether the financial crisis had an effect in changing the causal relationship between the two variables.

To empirically test the causal relationship between public debt and economic growth in a panel framework, we used the Dumitrescu and Hurlin (2012) Granger non-causality test which allows to conclude whether there is no causal relationship or if the causal relationship holds for at least one country. However, as it was previously referred, this alternative hypothesis presents a drawback because it does not allow to infer for which countries the causal relationship holds. Therefore, we conducted country-specific regressions to address this drawback.

Following the model specification of Ferreira (2016) and Gómez-Puig and Sosvilla-Rivero (2015), we applied the referred test on equations (6) and (7) for three different time frames, more specifically *i*) from 1995 to 2016, the full sample period, *ii*) from 1995 to 2006, the pre-crisis period, and *iii*) from 2007 to 2016, the post-crisis period:

$$\begin{aligned} \Delta GDP_{i,t} = & \beta_1 \Delta GDP_{i,t-1} + \beta_2 \Delta GDP_{i,t-2} + \beta_3 \Delta DebttoGDP_{i,t-1} \\ & + \beta_4 \Delta DebttoGDP_{i,t-2} + \mu_{i,t} \end{aligned} \quad (6)$$

$$\begin{aligned} \Delta DebttoGDP_{i,t} = & \beta_1 \Delta GDP_{i,t-1} + \beta_2 \Delta GDP_{i,t-2} + \beta_3 \Delta DebttoGDP_{i,t-1} \\ & + \beta_4 \Delta DebttoGDP_{i,t-2} + \mu_{i,t} \end{aligned} \quad (7)$$

4.2.1 Full Sample (1995-2016) Causality Analysis

We start by studying the causal relationship between public debt and economic growth for the entire sample. By applying the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (6), we verified that the p-values of both test statistics were below the significance level of 1%, rejecting the null hypothesis, and therefore suggesting that public debt Granger-cause economic growth for at least one country (see Table 1).

Table 1. Public Debt \Rightarrow Economic Growth (1995-2016): Dumitrescu and Hurlin (2012) Granger non-Causality test

Lag order	2
Average Wald statistic	7,0988
Average Z statistic	8,0620 (P-value = 0,0000)
Average Z-tilde statistic	5,8953 (P-value = 0,0000)
Number of observations	200

However, as it was previously discussed, by rejecting the null hypothesis we cannot infer for which countries such causal relationship holds. To address this matter, we conducted country-specific regressions¹¹ and we found evidence that a causal relationship departing from public debt to economic growth only holds for Belgium, Ireland, Luxembourg and Netherlands (see Table 2).

¹¹see tables C.1 and C.2 in the appendix for details.

Table 2. **Public Debt \Rightarrow Economic Growth (1995-2016): Country-Specific Causality**

Country	Wald statistic	P-value	Causality
Belgium	12,651135	0,00178995	YES
France	2,6337117	0,26797653	NO
Germany	2,5889329	0,27404405	NO
Greece	2,0645546	0,35619487	NO
Ireland	6,0447261	0,04868603	YES
Italy	0,07532267	0,96303903	NO
Luxembourg	10,39591	0,00552786	YES
Netherlands	30,218848	2,742E-07	YES
Portugal	4,1357988	0,12645113	NO
Spain	0,17951886	0,91415108	NO

By applying the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (7) to analyse the reverse causality, our results suggest that economic growth did not Granger-cause public debt, given that the p-values of both test statistics were above the significance level of 10%, therefore not rejecting the null hypothesis (see Table 3).

Table 3. **Economic Growth \Rightarrow Public Debt (1995-2016): Dumitrescu and Hurlin (2012) Granger non-Causality test**

Lag order	2
Average Wald statistic	2,4599
Average Z statistic	0,7272 (P-value = 0,4671)
Average Z-tilde statistic	0,2357 (P-value = 0,8136)
Number of observations	200

On the other hand, our country-specific regressions¹² provided evidence that this causal relationship holds for Germany. However, for the majority of the countries considered, this causal relationship did not hold (see Table 4).

¹²see tables C.1 and C.2 in the appendix for details.

Table 4. **Economic Growth \Rightarrow Public Debt (1995-2016): Country-Specific Causality**

Country	Wald-statistic	P-value	Causality
Belgium	4,0339089	0,13306009	NO
France	1,0376354	0,59522386	NO
Germany	12,658068	0,00178376	YES
Greece	1,6450296	0,43932545	NO
Ireland	0,14286936	0,93105709	NO
Italy	0,51636914	0,77245265	NO
Luxembourg	2,4792072	0,28949895	NO
Netherlands	1,3335599	0,51335897	NO
Portugal	0,38407523	0,82527583	NO
Spain	0,36831444	0,83180502	NO

In summary, these results suggest that, from 1995 to 2016, there was predominantly a one-way causal relationship departing from public debt to economic growth. Our country-specific regressions for this period suggest that in Belgium, Ireland, Luxembourg and Netherlands public debt Granger-cause economic growth while in Germany economic growth Granger-cause public debt.

4.2.2 Pre-Crisis (1995-2006) Causality Analysis

We now focus our analysis on the period from 1995 to 2006, which represents the pre-crisis period.

By applying the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (6), we verified that the p-values of both test statistics were above the significance level of 1%, not rejecting the null hypothesis, and therefore suggesting that public debt did not Granger-cause economic growth in our panel during the pre-crisis period. However, the average Z-statistic may provide ambiguous conclusions depending on the significance level considered (see table 5).

Table 5. **Public Debt \Rightarrow Economic Growth (1995-2006): Dumitrescu and Hurlin (2012) Granger non-Causality test**

Lag order	2
Average Wald-statistic	3,2415
Average Z statistic	1,9630 (P-value = 0,0496)
Average Z-tilde statistic	0,3264 (P-value = 0,7441)
Number of observations	100

Therefore, we conducted the country-specific regressions¹³ which allowed to assess that this causal relationship only hold for Germany and Ireland during the pre-crisis period (see Table 6).

Table 6. **Public Debt \Rightarrow Economic Growth (1995-2006): Country-Specific Causality**

Country	Wald-statistic	P-value	Causality
Belgium	1,0038704	0,60535804	NO
France	1,7402144	0,41890663	NO
Germany	9,5202509	0,00856454	YES
Greece	2,3681054	0,30603594	NO
Ireland	5,6680037	0,05877716	YES
Italy	0,6932546	0,7070688	NO
Luxembourg	2,7394471	0,25417722	NO
Netherlands	4,0892635	0,12942784	NO
Portugal	1,4754192	0,47820794	NO
Spain	3,1172494	0,21042527	NO

To infer about the reverse causality during the pre-crisis period, we applied the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (7) and verified that the p-values of both test statistics were below the significance level of 1%, rejecting the null hypothesis, and therefore suggesting that economic growth Granger-cause public debt on at least one country (see Table 7).

¹³see tables C.3 and C.4 in the appendix for details.

Table 7. **Economic Growth \Rightarrow Public Debt (1995-2006): Dumitrescu and Hurlin (2012) Granger non-Causality test**

Lag order	2
Average Wald-statistic	9,1376
Average Z statistic	11,2856 (P-value = 0,0000)
Average Z-tilde statistic	4,6858 (P-value = 0,0000)
Number of observations	100

To assess for which countries such causal relationship holds, we conducted country-specific regressions¹⁴ which allowed to conclude that economic growth Granger-cause public debt in Germany, Greece, Ireland, Italy and Netherlands (see Table 8).

Table 8. **Economic Growth \Rightarrow Public Debt (1995-2006): Country-Specific Causality**

Country	Wald-statistic	P-value	Causality
Belgium	0,54653446	0,76088943	NO
France	0,78192901	0,67640417	NO
Germany	8,6987519	0,01291487	YES
Greece	4,8653124	0,0878033	YES
Ireland	38,611765	4,126E-09	YES
Italy	4,8099862	0,09026612	YES
Luxembourg	0,38585835	0,82454037	NO
Netherlands	29,725766	3,509E-07	YES
Portugal	2,9162274	0,23267475	NO
Spain	0,03399233	0,98314746	NO

In summary, the analysis to the pre-crisis period, from 1995 to 2006, provided evidence that the causal relationship departed predominantly from economic growth to public debt. Additionally, the country-specific regressions allowed to conclude that, during this period, such causal relationship holds for Greece, Italy and Netherlands, while Germany and Ireland presented a bi-directional causal relationship between public debt and economic growth.

¹⁴see tables C.3 and C.4 in the appendix for details.

4.2.3 Post-Crisis (2007-2016) Causality Analysis

In this section we turn our focus to the period from 2007 to 2016, which represents the post-crisis period.

By applying the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (6) the conclusions drawn were ambiguous, given that the results lead to different conclusions depending on the test statistic considered, *i.e.* the average Z statistic rejected the null hypothesis while the average Z-tilde statistic did not (see Table 9).

Table 9. **Public Debt \Rightarrow Economic Growth (2007-2016): Dumitrescu and Hurlin (2012) Granger non-Causality test**

Lag order	2
Average Wald-statistic	7,1778
Average Z statistic	8,1869 (P-value = 0,0000)
Average Z-tilde statistic	1,6311 (P-value = 0,1029)
Number of observations	80

Therefore, we conducted country-specific regressions¹⁵ which showed that public debt Granger-cause economic growth in Greece, Ireland, Luxembourg, Netherlands and Portugal (see Table 10).

Table 10. **Public Debt \Rightarrow Economic Growth (2007-2016): Country-Specific Causality**

Country	Wald-statistic	P-value	Causality
Belgium	2,5213403	0,283464	NO
France	0,15031864	0,92759569	NO
Germany	3,5321676	0,17100135	NO
Greece	6,5472299	0,03786928	YES
Ireland	4,6363872	0,09845127	YES
Italy	3,9451615	0,13909742	NO
Luxembourg	16,91147	0,00021268	YES
Netherlands	22,886931	0,00001072	YES
Portugal	10,115551	0,00635969	YES
Spain	0,53182792	0,76650508	NO

¹⁵see tables C.5 and C.6 in the appendix for details.

Regarding the reverse causality, we applied the Dumitrescu and Hurlin (2012) Granger non-causality test on equation (7) and, similarly to the previous results, the conclusions drawn were ambiguous, provided that, depending on the test statistic considered we would arrive to different conclusions (see Table 11).

Table 11. **Economic Growth \Rightarrow Public Debt (2007-2016): Dumitrescu and Hurlin (2012) Granger non-Causality test**

Lag order	2
Average Wald-statistic	3,8679
Average Z statistic	2,9535 (P-value = 0,0031)
Average Z-tilde statistic	0,2268 (P-value = 0,8206)
Number of observations	80

Nevertheless, our country-specific regressions¹⁶ showed that economic growth Granger-cause public debt in Germany, Ireland and Luxembourg (see Table 12).

Table 12. **Economic Growth \Rightarrow Public Debt (2007-2016): Country-Specific Causality**

Country	Wald statistic	P-value	Causality
Belgium	2,8545071	0,23996708	NO
France	0,72231664	0,69686866	NO
Germany	5,8117327	0,05470138	YES
Greece	0,78431513	0,67559765	NO
Ireland	13,280428	0,00130675	YES
Italy	2,3247165	0,31274777	NO
Luxembourg	7,1834639	0,02755057	YES
Netherlands	1,4393504	0,48691037	NO
Portugal	3,2220595	0,19968189	NO
Spain	1,0565125	0,58963225	NO

In summary, the analysis to the post-crisis period, from 2007 to 2016, provided evidence that the causal relationship departed predominantly from public debt to economic growth. Additionally, the country-specific regressions allowed to conclude that, during this period,

¹⁶see tables C.5 and C.6 in the appendix for details.

such causal relationship hold for Greece, Netherlands and Portugal, while Ireland and Luxembourg presented a bi-directional causal relationship between public debt and economic growth. Similarly to the pre-crisis period, Germany presented a causal relationship departing from economic growth to public debt.

4.2.4 Robustness Checks

We also conducted robustness checks to verify whether our conclusions were robust to different methodologies. In this regard, our robustness checks were conducted by using different estimators, such as the Ordinary Least Squares (OLS) estimator, the Fixed Effects (FE) estimator to account for country heterogeneity, and the Generalized Method of Moments (GMM) estimator to account for potential endogeneity (see Tables 13 and 14).

Table 13. **Robustness Checks: Robust OLS and Fixed Effects**

	Robust OLS		Fixed Effects	
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$
$\Delta GDP_{i,t-1}$	0,456** (0,196)	-0,311 (0,432)	0,277*** (0,0878)	-0,414 (0,363)
$\Delta GDP_{i,t-2}$	0,087 (0,118)	0,524* (0,298)	-0,0538 (0,0899)	0,445 (0,372)
$\Delta DebttoGDP_{i,t-1}$	-0,0775*** (0,0153)	0,322** (0,140)	-0,107*** (0,0203)	0,299*** (0,0841)
$\Delta DebttoGDP_{i,t-2}$	0,0587*** (0,0183)	0,208*** (0,0749)	0,0179 (0,0214)	0,177** (0,0855)
Number of Observations	200	200	200	200
R^2	0,376	0,206	0,322	0,188

Standard Errors in parentheses

* $p < 0,1$; ** $p < 0,05$; *** $p < 0,01$

Table 14. **Robustness Checks: Generalized Method of Moments**

	2 Step GMM	
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$
ΔGDP_{IV}		-1,382 (1,132)
$\Delta GDP_{i,t-1}$	0,187* (0,102)	
$\Delta GDP_{i,t-2}$	-0,0585 (0,0794)	
$\Delta DebttoGDP_{IV}$	-0,242*** (0,0572)	
$\Delta DebttoGDP_{i,t-1}$		0,142 (0,170)
$\Delta DebttoGDP_{i,t-2}$		0,155** (0,0616)
Number of Observations	200	200
R^2	0,321	0,38

Standard Errors in parentheses

*p<0,1 ; **p<0,05; ***p<0,01

These regressions were conducted for the entire sample period, from 1995 to 2016, to take advantage from the larger number of observations and the conclusions drawn from these regressions support a causal relationship going from public debt to economic growth. More specifically, we verified that the lagged debt-to-GDP ratio annual growth rate was statistically significant in explaining the real GDP annual growth rate, while the contrary was not verified. Therefore, by following the causality definition of Granger (1969), these robustness checks support the idea of a one-way causality, *i.e.* public debt Granger-cause economic growth. Additionally, these regressions allowed to infer that public debt negatively affects economic growth, as the coefficients related with public debt showed that the overall effect was negative for the three estimators considered.

4.3 Public Debt Thresholds

After finding evidence that public debt predominantly causes economic growth in our sample, we conducted a threshold analysis using the Hansen (2000) endogenous threshold estimation

procedure as our main methodology and, to better support our results, we followed Bilan and Ihnatov (2015) and Afonso and Alves (2014) methodology, by using a partial derivative approach to estimate the maximum public debt threshold of our econometric model as a robustness check.

4.3.1 Threshold Estimation

$$\begin{aligned} \Delta GDP_{i,t} = & \beta_0 + \beta_1 DebttoGDP_{i,t} + \beta_2 DebttoGDP_{i,t}^2 + \beta_3 \Delta Population_{i,t} \\ & + \beta_4 TradeOpenessGDP_{i,t} + \beta_5 CapitalStockGDP_{i,t} \\ & + \beta_6 lnGDPperCapita_{i,t} + \beta_7 BudgetBalancePrimaryGDP_{i,t} \\ & + \beta_8 DebtServiceGDP_{i,t} + \mu_{i,t} \end{aligned} \quad (8)$$

Table 15. **Hansen (2000) Threshold Estimation**

Threshold Estimate	85,67%
95% Confidence Interval	[72,43%-85,67%]
Sum of Squared Errors	1337,1337
Residual Variance	6,6195
Joint R^2	0,4767
Heteroskedasticity Test (P-value)	0,3722
Number of observations	220

The results from the Hansen (2000) endogenous threshold estimation procedure provided a maximum public debt threshold of 85,67% with a 95% confidence interval for the threshold between 72,43% and 85,67%. The aforementioned procedure was applied to equation (8) and yielded the output above (see Table 15).

4.3.2 Robustness Checks

Our robustness checks, based on Bilan and Ihnatov (2015) and Afonso and Alves (2014) methodology, provided different results. After estimating a statistically significant model for each of the estimators¹⁷, we used a partial derivative approach to estimate the maximum public debt threshold (see equations (9) and (10)).

$$\frac{\partial \Delta GDP_{i,t}}{\partial DebttoGDP_{i,t}} = \beta_1 + 2\beta_2 DebttoGDP_{i,t} = 0 \quad (9)$$

¹⁷see tables from D.1 to D.4 in the appendix for details.

$$\Leftrightarrow DebttoGDP_{max} = \frac{-\beta_1}{2\beta_2} \quad (10)$$

Table 16. **Robustness Checks: Partial Derivative Approach**

	OLS	LSDV	FE	GMM
	$\Delta GDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta GDP_{i,t}$
$DebttoGDP_{i,t}$	0,05145** (0,0227)	0,0729612** (0,0319)	0,0729612** (0,0319)	0,0741053** (0,0319)
$DebttoGDP_{i,t}^2$	-0,0004158*** (0,000122)	-0,0006508*** (0,000148)	-0,0006508*** (0,000148)	-0,0006557*** (0,000148)
Threshold	61,87%	56,06%	56,06%	56,51%
Number of Observations	220	220	220	220
R^2	0,297	0,626	0,360	0,360

Standard errors in parentheses

*p<0,1; **p<0,05; ***p<0,01

By analysing Table 16, we verify that, depending on the estimator used, we get maximum public debt thresholds between 56,06% and 61,87%. The most efficient and consistent estimator used in this approach, the Generalized Method of Moments (GMM) estimator, provided a maximum public debt threshold of 56,51%.

In summary, our results support a Laffer-curve type relationship between public debt and economic growth, provided that the coefficients for the public debt and public debt squared terms were positive and negative, respectively, throughout all specifications. Therefore, with this kind of relationship between our variables of interest we were able to estimate the maximum value of the debt-to-GDP ratio, after which economic growth would be negatively affected by additional public debt increments.

Even though our analyses showed that the threshold results obtained are highly dependent on the chosen methodology, we were able to find evidence that a debt threshold of 90% debt-to-GDP ratio did not hold for our sample. Specifically, with our main methodology, the Hansen (2000) endogenous threshold estimation procedure, we found evidence of a maximum public debt threshold to be between 72,43% and 85,67% debt-to-GDP ratios. On the other hand, the most efficient and consistent estimator from our robustness checks provided evidence of a maximum public debt threshold of 56,51% debt-to-GDP ratio.

4.4 Summary of Results

From our replication exercise of Reinhart and Rogoff (2010) research, we concluded that a 90% public debt threshold did not hold for our dataset. Nevertheless, the negative correlation between public debt and economic growth, referred by most of the literature, was verified. We also enhanced our replication exercise by including two additional debt-to-GDP ratio ranges that allowed to conclude that we have average negative economic growth for debt-to-GDP ratios above 120% instead.

Regarding the debate of whether public debt causes economic growth or economic growth causes public debt, our results for the entire sample period suggest a one-way causal relationship departing from public debt to economic growth. Our country-specific regressions showed that the countries that present a causal relationship departing from public debt to economic growth were Belgium, Ireland, Luxembourg and Netherlands, while for the reverse causality, *i.e.* departing from economic growth to public debt, only Germany presented such causal relationship (see Table 17). The robustness checks conducted for this time-frame also supported a one-way causality departing from public debt to economic growth (see Tables 13 and 14).

Table 17. **Causality Results Full Sample: Summary**

	Full sample (1995-2016)	
	Debt \Rightarrow Growth	Growth \Rightarrow Debt
Belgium	YES	NO
France	NO	NO
Germany	NO	YES
Greece	NO	NO
Ireland	YES	NO
Italy	NO	NO
Luxembourg	YES	NO
Netherlands	YES	NO
Portugal	NO	NO
Spain	NO	NO

The results for the pre-crisis period suggest a one-way causality departing from economic growth to public debt. Our country-specific regressions showed that such causal relationship only hold for Greece, Italy and Netherlands while for Germany and Ireland we found evidence of a bi-directional causal relationship. A possible explanation for economic growth causing

public debt may be due to the automatic stabilizers (Panizza and Presbitero, 2014; Gómez-Puig and Sosvilla-Rivero, 2015), whose effect might be more pronounced in countries with large welfare systems.

On the other hand, for the post-crisis period, the test statistics from our main test lead us to ambiguous conclusions. However, by analysing the country-specific regressions, we found evidence supporting a causal relationship departing from public debt to economic growth for Greece, Netherlands and Portugal, a bi-directional causal relationship for Ireland and Luxembourg and a causal relationship departing from economic growth to public debt for Germany (see Table 18). A possible explanation for public debt causing economic growth may be due to the creditor’s negative expectations about a country’s solvency, *i.e.* the debt overhang phenomenon discussed by Krugman (1988), which may lead the level of indebtedness to an unsustainable level due to higher interest rates.

Table 18. **Causality Results Pre-Crisis and Post-Crisis: Summary**

	Pre-Crisis (1995-2006)		Post-Crisis (2007-2016)	
	Debt⇒Growth	Growth⇒Debt	Debt⇒Growth	Growth⇒Debt
Belgium	NO	NO	NO	NO
France	NO	NO	NO	NO
Germany	YES	YES	NO	YES
Greece	NO	YES	YES	NO
Ireland	YES	YES	YES	YES
Italy	NO	YES	NO	NO
Luxembourg	NO	NO	YES	YES
Netherlands	NO	YES	YES	NO
Portugal	NO	NO	YES	NO
Spain	NO	NO	NO	NO

By comparing the pre-crisis and post-crisis results, we found evidence that the global financial crisis had an effect in changing the causal relationship between public debt and economic growth, given that the predominant causal relationship was from economic growth to public debt and from public debt to economic growth during the pre-crisis and post-crisis periods, respectively. Therefore, this suggests that the international economic growth context may have an impact in the causal relationship between public debt and economic growth. Additionally, our country-specific regressions, similarly to Donayre and Taivan (2017) and Gómez-Puig and Sosvilla-Rivero (2015), provided evidence that the causal relationship be-

tween public debt and economic growth varies across countries (see Tables 17 and 18).

Finally, our public debt threshold analysis pointed towards a maximum public debt threshold of 85,67%. However, our robustness checks provided different results depending on the estimator used. Considering the most efficient and consistent estimator used, the Generalized Method of Moments (GMM) estimator, we got a maximum public debt threshold of 56,51% (see Table 19).

Table 19. **Public Debt Thresholds: Summary**

Methodology	Threshold
Hansen (2000)	85,67%
OLS	61,87%
LSDV	56,06%
FE	56,06%
GMM	56,51%

5 Policy Implications

In terms of policy implications, our research's results imply that the theoretical and empirical research that initially supported austerity policies for highly indebted countries, namely Reinhart and Rogoff (2010) research, was not robust. Specifically, our replication results suggest that, by following Reinhart and Rogoff (2010) rationale, the EMU countries analysed would have a maximum public debt threshold of 120% debt-to-GDP ratio.

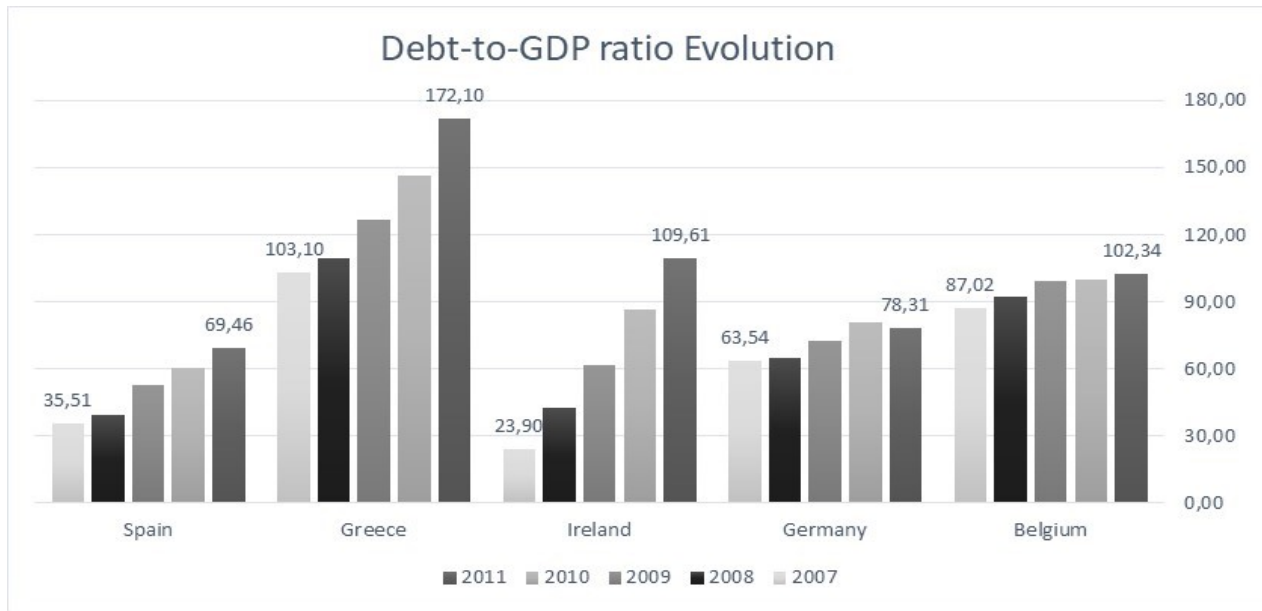


Figure 4: Public Debt evolution: Crisis' early stages (1/2). Source: AMECO

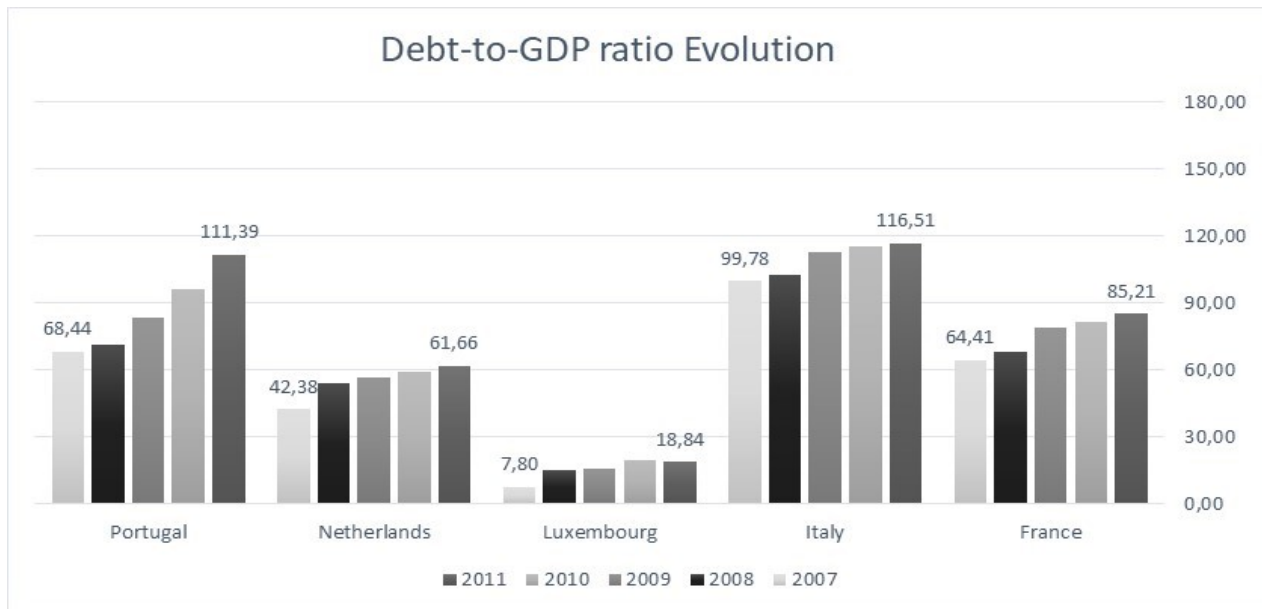


Figure 5: Public Debt evolution: Crisis' early stages (2/2). Source: AMECO

As we can see from the Figures 4 and 5, for the exception of Greece, none of the countries analysed presented debt-to-GDP ratios above such level at the early stages of the global financial crisis. Therefore, according to our replication results, governments, at the time, would still have enough fiscal space to pursue expansionary policies and would not need to apply severe austerity policies to decrease their debt burden. Nevertheless, the conclusions drawn from this methodology appear to be sensitive to the data and time-frame considered

which further supports the need to use other methodologies to infer about the maximum level of public debt that each economy can sustain.

Additionally, it was also found by Blanchard and Leigh (2013) research that the IMF's fiscal multipliers estimations were significantly optimistic. Therefore, cutting government expenditure would have a significantly larger negative impact on the economy's growth path, implying that the application of severe government expenditure cuts would not be appropriate. As a result, it might be the case that the implementation of severe austerity policies with a short temporal horizon was not the optimal path to pursue to remove the global economy out of the slump.

Such considerations do not imply that countries should disregard the stability of their public finances. When countries increase public debt levels by a significant amount, it leads the markets and creditors to demand a higher risk-premium, it crowds-out private investment and creates the expectation of possible future tax increases (Puente-Ajovín and Sanso-Navarro, 2015) which negatively affects the economy's growth path.

Our public debt threshold analysis showed that there is a Laffer-curve type relationship between public debt and economic growth. This means that additional public debt increments will present a negative impact on the economy's growth after surpassing a given threshold that we estimate to be between 72,43% and 85,67% debt-to-GDP ratios, which is far below the current debt level of most countries analysed in this research. Additionally, our results suggest that a balanced budget balance has a positive effect on economic growth, which further supports the importance of proper management of public finances (see Tables from D.1 to D.4 in the appendix).

However, we cannot claim that high levels of public debt will severely impair economic growth in all countries, given that our results provided evidence of significant heterogeneity in the causal relationship between public debt and economic growth among the EMU countries analysed. Therefore, we believe that different countries may present different public debt thresholds which makes the application of homogeneous fiscal policies throughout the EMU countries inappropriate.

As a result, considering our analysis, the 60% debt-to-GDP ratio limit inscribed in the European Stability-Growth Pact provides a safe benchmark for the countries to pursue because it would allow the countries to comply with the treaty's terms, to signal the markets and creditors that a proper management of public resources is being made and it would also endow countries with additional fiscal space to accommodate future unexpected negative economic shocks without risking surpassing our estimated maximum public debt threshold. Therefore, given that highly indebted countries currently have debt-to-GDP ratios significantly above

our threshold estimates, in the near future such countries are advised and expected to significantly reduce or to impose a downward trajectory on their debt burden, given that more than the sheer amount of government debt, it is the debt dynamic that seems more relevant for economic growth (Pescatori et al., 2014).

That being said, in the current context of the countries' public finances, the alternative for severe austerity policies with a short temporal horizon could be austerity policies with a longer temporal horizon. What we mean with this is that countries should conduct policies that guarantee a decrease of their debt burden while, at the same time, the implementation of such policies is diluted through time so that the stress caused in the economy is less pronounced. This alternative provides a number of challenges for the economies facing such endeavour, such as the need for the creditors to believe in the country's commitment and allow this alternative without deteriorating the financing conditions, and the need for within-country political coordination for the application of long-term measures, which might not be easy when the political cycles range between 4 to 8 years.

As it was previously referred, our results showed that the budget balance has a positive impact on economic growth. This variable includes, on the expenditure side, the debt service, *i.e.* the debt interests and capital paid by the government to their creditors, therefore public debt deviates public resources to accommodate this cost. However, to mitigate the impact of such cost, governments could try to renegotiate the debt terms with the creditors which would make the pressure of the debt service on the budget balance less pronounced and, consequently, allowing for a better economic performance. Supporting this suggestion we have Afonso and Jalles (2013) research that found evidence that longer debt maturities positively affected economic growth. Similarly to the previous suggestion, this one also presents a number of challenges and, even though such idea is being discussed to some extent, it is still not well accepted by the creditors.

6 Concluding Remarks

In this research, we analysed several topics regarding the short-run relationship between public debt and economic growth for a panel of 10 EMU countries between 1995 and 2016. First, we conducted a replication exercise of Reinhart and Rogoff (2010) to assess whether the 90% debt-to-GDP ratio threshold holds for our sample. Then, we analysed the causal relationship between public debt and economic growth, to fill the gap of most of the literature on the debt-growth nexus departing from the assumption of public debt causing economic growth without clearly testing such hypothesis. Specifically, we analysed the causal relationship for

the full sample and also by splitting it to analyse the effects of the global financial crisis that started in 2007. Finally, we endogenously determined the maximum public debt threshold for our sample using the Hansen (2000) endogenous threshold estimation procedure.

First, our replication exercise did not provide evidence of average negative growth after surpassing the 90% debt-to-GDP ratio. However, by enhancing this exercise, we found evidence of average negative growth for debt-to-GDP ratios above 120%. As a result, by following Reinhart and Rogoff (2010) rationale, we would be able to claim that there was a common and maximum public debt threshold at 120% debt-to-GDP ratio. However, considering the rigid assumptions of this methodology, we did not claim such result and conducted further analyses to obtain more meaningful results.

Second, the analysis of the causal relationship between public debt and economic growth provided interesting conclusions. When we focused our analysis on the entire sample period, from 1995 to 2016, our results suggest that there was predominantly a one-way causality departing from public debt to economic growth. Our robustness checks, also focused on this time-span, completely support this result throughout the methodologies used. However, when we compared the results of the pre-crisis (1995-2006) and post-crisis (2007-2016) analyses, we found evidence that the global financial crisis changed the causal relationship between public debt and economic growth, given that the predominant causal relationship was from economic growth to public debt and from public debt to economic growth during the pre-crisis and post-crisis periods, respectively. Additionally, our country-specific regressions provided evidence of a significant heterogeneity in the causal relationship between countries, suggesting that high levels of public debt may not cause low growth for all the countries analysed.

Finally, our threshold analysis provided evidence of a Laffer-curve (inverted U-shape curve) type relationship between public debt and economic growth. Such result implies that at lower levels of debt, public spending positively affects economic growth while after surpassing a certain threshold, which we estimate to be at 85,67% debt-to-GDP ratio, additional debt increments impair economic growth. Our robustness checks also support a Laffer-curve relationship however, it did not fully support the threshold results, given that the most efficient and consistent estimator used with the partial derivative approach yielded a maximum public debt threshold of 56,51%. Nonetheless, the results provided by this analysis are to be taken with caution, given that we found evidence of significant heterogeneity in the debt-growth relationship across countries. As a result, it is possible that each country has a specific public debt threshold or no threshold at all.

Regarding the policy implications, our research indicates that the austerity policies, implemented throughout the EMU countries in the aftermath of the financial crisis, found

empirical support on results that were not robust. Nevertheless, our results indicate that governments in most of the countries analysed should pursue policies to reduce or to impose a downward trajectory on their debt burden to improve economic performance.

In terms of further research, there are several possible extensions for this investigation. Given the heterogeneity found in our analysis regarding the causal relationship between public debt and economic growth, the first possible extension is to conduct country-specific regressions to estimate the maximum public debt threshold for each country. This would allow the use of specific control variables that could vary depending on the country under analysis.

Another possible extension is to analyse whether there are significant differences in the debt-growth relationship between the peripheral and central EMU countries, given that the former suffered more with the global financial crisis than the latter.

Finally, a third possible extension is to analyse the channels through which public debt affects economic growth.

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Appendix

A Variables List

- **BudgetBalanceGDP:** Budget Balance (Total Government Revenue (URTG) - Total Government Expenditure excluding (UUTG)) as percentage of Nominal Gross Domestic Product (UVGD) (**Source:** AMECO¹⁸);
- **BudgetBalancePrimaryGDP:** Budget Balance (Total Government Revenue (URTG) - Total Government Expenditure excluding interests (UUTGI)) as percentage of Nominal Gross Domestic Product (UVGD) (**Source:** AMECO);
- **CapitalStockGDP:** Real Gross Fixed Capital Stock (OIGT) as percentage of Real Gross Domestic Product (OVGD) (**Source:** AMECO);
- **DebtServiceGDP:** Interests paid by the government (UYIG) as percentage of Nominal Gross Domestic Product (UVGD) (**Source:** AMECO);
- **DebttoGDP:** Consolidated Government Gross Debt as percentage of Gross Domestic Product (UDGG) (**Source:** AMECO);
- Δ **GDP:** Annual growth rate of Real Gross Domestic Product (OVGD) (**Source:** AMECO);
- **TradeOpennessGDP:** Measures a country's openness to foreign economies (Real Imports of Goods and Services (OMGS) + Real Exports of Goods and Services (OXGS)) as percentage of Real Gross Domestic Product (OVGD) (**Source:** AMECO);
- **lnGDPperCapita:** logarithmic value of Real Gross Domestic Product *per capita* (RVGDP) (**Source:** AMECO);
- Δ **Population:** Annual growth rate of Total Population (NPTD) (**Source:** AMECO).

¹⁸AMECO variables' codes in parentheses.

B Reinhart and Rogoff (2010) Replication

Table B.1. Reinhart and Rogoff (2010) Replication

Debt-to-GDP Ratio	Average Growth	Median Growth
< 30%	4,00%	4,19%
30%-60%	2,72%	2,31%
60%-90%	2,33%	1,96%
+90%	1,03%	1,35%

Table B.2. Reinhart and Rogoff (2010) Enhanced Replication

Debt-to-GDP Ratio	Average Growth	Median Growth
< 30%	4,00%	4,19%
30%-60%	2,72%	2,31%
60%-90%	2,33%	1,96%
90%-120%	1,77%	1,65%
+120%	-1,22%	0,09%

C Causality

Table C.1. Full Sample (1995-2016): Country-Specific Regressions (1/2)

	Belgium	France	Germany	Greece	Ireland
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$
$\Delta GDP_{i,t-1}$	-0,6812* (0,3205)	0,4838 (0,5091)	-0,3622 (0,3123)	0,6793* (0,3255)	-1,0497 (0,8892)
$\Delta GDP_{i,t-2}$	-0,1538 (0,3567)	0,1859 (3,6968)	-0,5417* (0,2968)	-0,1229 (0,2576)	0,4777 (0,7037)
$\Delta DebttoGDP_{i,t-1}$	-0,4569*** (0,1419)	0,0458 (0,1821)	-0,1755 (0,1689)	-0,1475 (0,1212)	-0,0487 (0,3309)
$\Delta DebttoGDP_{i,t-2}$	0,0338 (0,1778)	0,1388 (0,1254)	-0,4237 (0,3595)	-0,1008 (0,1217)	-0,0939 (0,3323)
Constant	2,8426*** (0,8553)	-2,7398 (2,0191)	2,5001** (0,9502)	1,0754 (0,9821)	4,0292 (2,6825)
Number of observations	20	20	20	20	20
R^2	0,4362	0,2670	0,3122	0,7219	0,1473

Standard Errors in parentheses

*p<0,1 ; **p<0,05; ***p<0,01

Table C.2. Full Sample (1995-2016): Country-Specific Regressions (2/2)

	Italy	Luxembourg	Netherlands	Portugal	Spain
$\Delta GDP_{i,t}$	0,3361 (0,5457)	-0,0496 (0,3162)	0,6405** (0,2302)	0,1821 (0,3086)	0,7393 (0,6516)
$\Delta DebttoGDP_{i,t}$	-0,5555 (0,8308)	3,1359 (2,5103)	-1,6773 (1,6757)	0,2305 (0,8349)	0,1009 (3,1868)
$\Delta GDP_{i,t-1}$	-0,0683 (0,3826)	-0,1458 (0,2072)	-0,1197 (0,1527)	-0,1439 (0,3007)	0,0116 (0,5925)
$\Delta GDP_{i,t-2}$	0,0056 (0,3613)	0,2032 (0,2945)	0,1333 (0,2813)	-0,1348 (0,1122)	-0,0347 (0,1329)
$\Delta DebttoGDP_{i,t-1}$	0,1843 (0,5502)	-0,1052** (0,0371)	-0,1542*** (0,0386)	0,9849*** (0,3035)	0,8193 (0,6501)
$\Delta DebttoGDP_{i,t-2}$	-0,0678 (0,3548)	-0,01509 (0,0464)	0,1273** (0,0531)	-0,4640 (0,8134)	-0,4374 (0,5902)
Constant	0,4887 (0,7793)	5,4822** (1,9969)	0,8312 (0,5517)	2,3180 (1,0301)	0,4898 (1,4885)
Number of observations	20	20	20	20	20
R^2	0,1025	0,4516	0,7388	0,4703	0,5721

Standard Errors in parentheses
 *p<0,1 ; **p<0,05; ***p<0,01

Table C.3. Pre-Crisis (1995-2006): Country-Specific Regressions (1/2)

	Belgium	France	Germany	Greece	Ireland							
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$							
$\Delta GDP_{i,t-1}$	-0,4131 (0,5589)	1,1356 (1,0005)	-1,7305 (3,0361)	2,4678** (0,9375)	-4,4846* (1,7992)	-0,6548 (0,4148)	$\Delta DebttoGDP_{i,t}$	0,9631 (0,6066)	$\Delta GDP_{i,t}$	0,7008 (0,3555)	$\Delta DebttoGDP_{i,t}$	-2,7271*** (0,5207)
$\Delta GDP_{i,t-2}$	-0,2809 (0,5611)	0,3338 (0,5437)	0,2772 (0,5796)	-1,1951 (1,7588)	0,4289 (0,4546)	-0,6418 (0,8725)	-0,4034 (0,5247)	-0,5401 (0,7673)	1,4039* (0,6770)	-0,3695 (0,9916)		
$\Delta DebttoGDP_{i,t-1}$	-0,2097 (0,5387)	-0,4101 (0,5220)	0,3702 (0,4843)	-0,1559 (1,4696)	1,1669* (0,4894)	-1,6320 (0,9393)	-0,2789 (0,1876)	0,5599* (0,2743)	0,5009 (0,2512)	-0,7689* (0,3679)		
$\Delta DebttoGDP_{i,t-2}$	-0,4777 (0,5708)	-0,0138 (0,5531)	-0,0163 (0,1673)	-0,1835 (0,5077)	-0,1830 (0,1372)	0,3201 (0,2633)	0,1269 (0,1992)	-0,9862** (0,2914)	0,2797 (0,1866)	0,0993 (0,2734)		
Constant	1,8866 (2,1107)	-5,5624** (2,0453)	-1,5215 (3,7053)	8,4019 (11,2437)	-3,9245 (2,1422)	10,2262* (4,1111)	8,2084** (2,9452)	-0,9865 (4,3071)	-1,3173 (2,8909)	6,9049 (4,2341)		
Number of observations	10	10	10	10	10	10	10	10	10	10		
R^2	0,1803	0,2548	0,3348	0,3708	0,6204	0,6438	0,4984	0,7866	0,6617	0,8678		

Standard Errors in parentheses
 *p<0,1 ; **p<0,05; ***p<0,01

Table C.4. Pre-Crisis (1995-2006): Country-Specific Regressions (2/2)

	Italy	Luxembourg	Netherlands	Portugal	Spain				
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$				
$\Delta GDP_{i,t-1}$	0,3735 (0,4742)	0,1282 (0,5938)	-0,9582 (1,8483)	1,1094** (0,4188)	-1,1216 (0,9716)	0,3669 (0,5699)	-1,7428 (1,7278)	0,9263** (0,3346)	-0,0338 (0,9386)
$\Delta GDP_{i,t-2}$	-0,2784 (0,4019)	-1,0844 (0,5864)	-0,1156 (0,3598)	0,2987 (1,1198)	0,3238 (0,7535)	-4,1775* (1,7479)	-0,0247 (0,5190)	-1,4743 (1,5736)	-0,5564 (0,3245)
$\Delta DebttoGDP_{i,t-1}$	0,1779 (0,2796)	0,3439 (0,4080)	-0,2663 (0,2044)	0,0825 (0,6363)	0,2905 (0,1815)	-0,9277* (0,4210)	-0,1769 (0,1750)	0,2512 (0,5307)	0,1525 (0,1062)
$\Delta DebttoGDP_{i,t-2}$	0,0517 (0,1888)	0,0858 (0,2755)	-0,0297 (0,1036)	-0,1342 (0,3225)	0,1539 (0,1656)	-0,9185* (0,3842)	0,0339 (0,2038)	-0,8115 (0,6179)	-0,0867 (0,2362)
Constant	1,7554 (0,9536)	0,5327 (1,3914)	4,9233 (2,9140)	3,4532 (9,0703)	0,4358 (0,8254)	3,8895* (1,9148)	1,6850 (2,2402)	10,2451 (6,7919)	2,8178** (1,0172)
Number of observations	10	10	10	10	10	10	10	10	10
R^2	0,1655	0,4240	0,4938	0,1514	0,8433	0,8738	0,5921	0,5560	0,6342

Standard Errors in parentheses
 *p<0,1 ; **p<0,05; ***p<0,01

Table C.5. Post-Crisis (2007-2016): Country-Specific Regressions (1/2)

	Belgium		France		Germany		Greece		Ireland	
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$
$\Delta GDP_{i,t-1}$	-0,4948 (0,4122)	1,0859 (0,9208)	-0,4387 (1,1968)	2,3266 (4,6855)	-0,8283 (0,4968)	-0,4242 (1,1839)	0,5257 (0,4153)	-0,7339 (2,8093)	-0,0804 (0,7294)	1,2871 (0,6078)
$\Delta GDP_{i,t-2}$	-0,3036 (0,4564)	-0,2016 (1,0197)	-0,3450 (1,1209)	0,4826 (4,3884)	-0,8018 (0,4621)	1,6183 (1,1011)	-0,4237 (0,2674)	1,1614 (1,8088)	-1,3829 (1,6199)	0,4148 (1,3500)
$\Delta DebttoGDP_{i,t-1}$	-0,1880 (0,2100)	0,7134 (0,4692)	-0,0798 (0,3946)	0,8429 (1,5447)	-0,2800 (0,2447)	0,7615 (0,5830)	-0,1626 (0,1060)	-0,1200 (0,7174)	-0,1861 (0,1782)	0,8688** (0,1485)
$\Delta DebttoGDP_{i,t-2}$	0,1733 (0,2129)	-0,3095 (0,4756)	0,1304 (0,4361)	-0,3002 (1,7073)	0,2517 (0,1945)	0,0791 (0,4635)	-0,1966 (0,1121)	0,0813 (0,7582)	-0,1871 (0,2029)	0,3051 (0,1691)
Constant	1,6351 (1,2613)	0,0433 (2,8180)	0,8062 (2,5602)	0,2154 (10,0231)	2,7422 (1,4569)	-1,9641 (3,4718)	-0,3985 (1,2005)	8,1201 (8,1209)	13,9212 (7,4067)	-1,9511* (6,1725)
Number of observations	8	8	8	8	8	8	8	8	8	8
R^2	0,7956	0,6199	0,5054	0,2144	0,6567	0,6472	0,8719	0,1379	0,5856	0,9587

Standard Errors in parentheses

*p<0,1 ; **p<0,05; ***p<0,01

Table C.6. Post-Crisis (2007-2016): Country-Specific Regressions (2/2)

	Italy	Luxembourg	Netherlands	Portugal	Spain					
	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$	$\Delta DebttoGDP_{i,t}$	$\Delta GDP_{i,t}$					
$\Delta GDP_{i,t-1}$	-2,4516 (2,0852)	3,3989 (2,9099)	-0,4985 (0,4248)	-1,9873 (1,3453)	-0,2805 (0,4663)	1,5913 (2,3120)	-0,2845 (0,3811)	1,5858 (1,2289)	0,7557 (1,8156)	-1,4359 (7,9818)
$\Delta GDP_{i,t-2}$	-0,1966 (0,9005)	0,3750 (1,2567)	-0,4162 (0,2286)	0,3891 (0,7239)	-0,2660 (0,1841)	-0,4480 (0,9125)	-0,9667 (0,4344)	0,7717 (1,4008)	-1,4315 (1,9689)	6,4091 (8,6559)
$\Delta DebttoGDP_{i,t-1}$	-1,4499 (1,4986)	2,3358 (2,0912)	-0,0949* (0,0316)	-0,0481 (0,1002)	-0,1607** (0,0467)	0,3794 (0,2315)	0,0269 (0,1486)	1,0941 (0,4792)	0,0839 (0,4177)	-0,1064 (1,8361)
$\Delta DebttoGDP_{i,t-2}$	0,8135 (0,5741)	-1,1059 (0,8011)	-0,0377 (0,0494)	0,1037 (0,1564)	-0,0254 (0,0973)	0,5227 (0,4825)	-0,3315 (0,1692)	-0,2931 (0,5456)	-0,1945 (0,3519)	0,7746 (1,5472)
Constant	-0,7734 (2,9445)	2,2058 (4,1090)	6,4516* (2,3222)	6,2561 (7,3532)	1,8024 (0,9493)	-3,8725 (4,7067)	1,3230 (1,1528)	2,7474 (3,7172)	1,0066 (4,7058)	5,8456 (20,6873)
Number of observations	8	8	8	8	8	8	8	8	8	8
R^2	0,7003	0,6890	0,8506	0,8382	0,9103	0,5155	0,7161	0,7488	0,3696	0,3972

Standard Errors in parentheses

*p<0,1 ; **p<0,05; ***p<0,01

D Public Debt Thresholds

Table D.1. Robustness Check: Ordinary Least Squares (OLS)

	$\Delta GDP_{i,t}$
Constant	12,1887*** (3,2742)
$DebttoGDP_{i,t}$	0,05145** (0,0227)
$DebttoGDP_{i,t}^2$	-0,0004158*** (0,0001)
$\ln.GDPperCapita_{i,t}$	-3,5418*** (1,0171)
$TradeOpennessGDP_{i,t}$	0,0191*** (0,0051)
$BudgetBalanceGDP_{i,t}$	0,3603*** (0,0585)
Number of observations	220
F(5, 214)	18,05
Prob > F	0,0000
R^2	0,2966
Adjusted R^2	0,2801
Root MSE	2,898

Standard errors in parentheses

*p<0,1; **p<0,05; ***p<0,01

Table D.2. Robustness Check: Least Squares Dummy Variables (LSDV)

	$\Delta GDP_{i,t}$
$DebttoGDP_{i,t}$	0,0729612** (0,0319)
$DebttoGDP_{i,t}^2$	-0,0006508*** (0,0001)
$\Delta Population_{i,t}$	-2,0361*** (0,5263)
$BudgetBalanceGDP_{i,t}$	0,4821*** (0,0567)
Dummy Variable for each country	
Number of observations	220
F(14, 206)	24,61
Prob > F	0,000
R^2	0,6259
Adjusted R^2	0,6004
Root MSE	2,5383

Standard errors in parentheses

*p<0,1; **p<0,05; ***p<0,01

Table D.3. **Robustness Check: Fixed Effects (FE)**

	$\Delta GDP_{i,t}$
Constant	3,8938** (1,6245)
$DebttoGDP_{i,t}$	0,0729612** (0,0319)
$DebttoGDP_{i,t}^2$	-0,0006508*** (0,0001)
$\Delta Population_{i,t}$	-2,0361*** (0,5263)
$BudgetBalanceGDP_{i,t}$	0,4820*** (0,0567)
Number of observations	220
F(4, 206)	28,98
Prob > F	0,0000
R^2 :	
Within	0,3601
Between	0.000
Overall	0,1622

Standard errors in parentheses

*p<0,1; **p<0,05; ***p<0,01

Table D.4. **Robustness Check: 2-Step Generalized Method of Moments (GMM)**

	$\Delta GDP_{i,t}$
$DebttoGDP_{i,t}$	0,0741053** (0,0319734)
$DebttoGDP_{i,t}^2$	-0,0006557*** (0,000148)
$\Delta Population_{i,t}$	-2,0299*** (0,5264)
$BudgetBalance_{i,t}$	0,4827*** (0,0567)
Number of observations	220
F(4, 206)	29,02
Prob > F	0,0000
Centered R^2	0,3601
Uncentered R^2	0,3601
Root MSE	2,538

Standard errors in parentheses

*p<0,1; **p<0,05; ***p<0,01